

Outline

Balls in Bins.

Outline

Balls in Bins.

Birthday.

Outline

Balls in Bins.

Birthday.

Coupon Collector.

Outline

Balls in Bins.

 Birthday.

 Coupon Collector.

 Load balancing.

Outline

Balls in Bins.

Birthday.

Coupon Collector.

Load balancing.

Outline

Balls in Bins.

 Birthday.

 Coupon Collector.

 Load balancing.

Geometric Distribution: Memoryless property.

Outline

Balls in Bins.

 Birthday.

 Coupon Collector.

 Load balancing.

Geometric Distribution: Memoryless property.

Poisson Distribution: Sum of two Poisson is Poisson.

pause

Outline

Balls in Bins.

 Birthday.

 Coupon Collector.

 Load balancing.

Geometric Distribution: Memoryless property.

Poisson Distribution: Sum of two Poisson is Poisson.

pause

Tail Sum for Expectation.

Outline

Balls in Bins.

 Birthday.

 Coupon Collector.

 Load balancing.

Geometric Distribution: Memoryless property.

Poisson Distribution: Sum of two Poisson is Poisson.

pause

Tail Sum for Expectation.

Regression (optional.)

Balls in bins

Balls in bins

One throws m balls into $n > m$ bins.

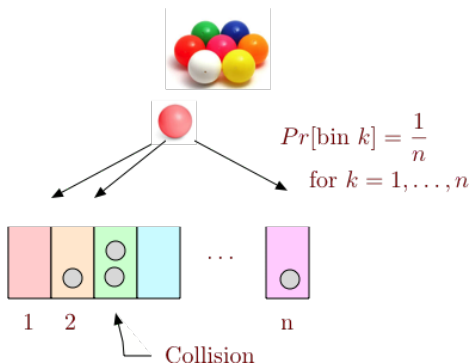
Balls in bins

One throws m balls into $n > m$ bins.



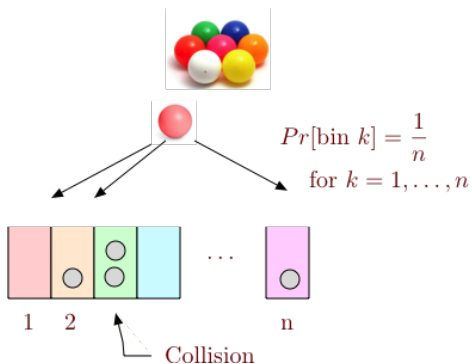
Balls in bins

One throws m balls into $n > m$ bins.



Balls in bins

One throws m balls into $n > m$ bins.



Theorem:

$Pr[\text{no collision}] \approx \exp\left\{-\frac{m^2}{2n}\right\}$, for large enough n .

Balls in bins

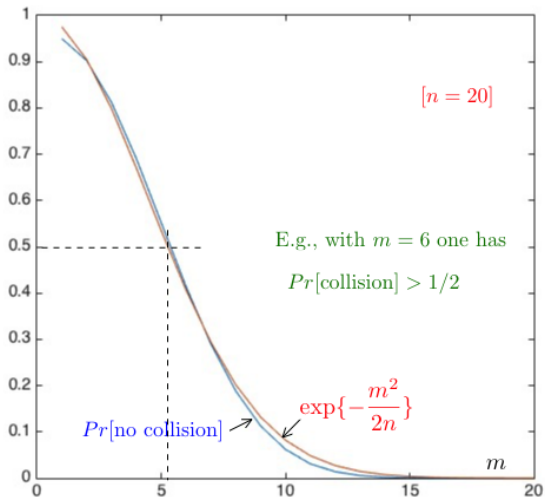
Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .

Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .



Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\left\{-\frac{m^2}{2n}\right\}$, for large enough n .

Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .

In particular, $Pr[\text{no collision}] \approx 1/2$ for $m^2/(2n) \approx \ln(2)$, i.e.,

$$m \approx \sqrt{2\ln(2)n} \approx 1.2\sqrt{n}.$$

Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .

In particular, $Pr[\text{no collision}] \approx 1/2$ for $m^2/(2n) \approx \ln(2)$, i.e.,

$$m \approx \sqrt{2\ln(2)n} \approx 1.2\sqrt{n}.$$

E.g., $1.2\sqrt{20} \approx 5.4$.

Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .

In particular, $Pr[\text{no collision}] \approx 1/2$ for $m^2/(2n) \approx \ln(2)$, i.e.,

$$m \approx \sqrt{2\ln(2)n} \approx 1.2\sqrt{n}.$$

E.g., $1.2\sqrt{20} \approx 5.4$.

Roughly, $Pr[\text{collision}] \approx 1/2$ for $m = \sqrt{n}$.

Balls in bins

Theorem:

$Pr[\text{no collision}] \approx \exp\{-\frac{m^2}{2n}\}$, for large enough n .

In particular, $Pr[\text{no collision}] \approx 1/2$ for $m^2/(2n) \approx \ln(2)$, i.e.,

$$m \approx \sqrt{2\ln(2)n} \approx 1.2\sqrt{n}.$$

E.g., $1.2\sqrt{20} \approx 5.4$.

Roughly, $Pr[\text{collision}] \approx 1/2$ for $m = \sqrt{n}$. ($e^{-0.5} \approx 0.6$.)

The Calculation.

A_i = no collision when i th ball is placed in a bin.

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \cdots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$Pr[A_i | A_{i-1} \cap \dots \cap A_1] = (1 - \frac{i-1}{n}).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$\Pr[A_1 \cap \dots \cap A_m] = \Pr[A_1] \Pr[A_2 | A_1] \dots \Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$Pr[A_1 \cap \dots \cap A_m] = Pr[A_1] Pr[A_2 | A_1] \dots Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$Pr[A_1 \cap \dots \cap A_m] = Pr[A_1] Pr[A_2 | A_1] \dots Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

Hence,

$$\ln(Pr[\text{no collision}]) = \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right)$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$Pr[A_1 \cap \dots \cap A_m] = Pr[A_1] Pr[A_2 | A_1] \dots Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

Hence,

$$\ln(Pr[\text{no collision}]) = \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right) \approx \sum_{k=1}^{m-1} \left(-\frac{k}{n}\right) (*)$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$\Pr[A_1 \cap \dots \cap A_m] = \Pr[A_1] \Pr[A_2 | A_1] \dots \Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow \Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

Hence,

$$\begin{aligned} \ln(\Pr[\text{no collision}]) &= \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right) \approx \sum_{k=1}^{m-1} \left(-\frac{k}{n}\right) (*) \\ &= -\frac{1}{n} \frac{m(m-1)}{2} (\dagger) \approx \end{aligned}$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$\Pr[A_1 \cap \dots \cap A_m] = \Pr[A_1] \Pr[A_2 | A_1] \dots \Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow \Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

Hence,

$$\begin{aligned} \ln(\Pr[\text{no collision}]) &= \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right) \approx \sum_{k=1}^{m-1} \left(-\frac{k}{n}\right) (*) \\ &= -\frac{1}{n} \frac{m(m-1)}{2} (\dagger) \approx -\frac{m^2}{2n} \end{aligned}$$

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$\Pr[A_1 \cap \dots \cap A_m] = \Pr[A_1] \Pr[A_2 | A_1] \dots \Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow \Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

Hence,

$$\begin{aligned} \ln(\Pr[\text{no collision}]) &= \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right) \approx \sum_{k=1}^{m-1} \left(-\frac{k}{n}\right) \quad (*) \\ &= -\frac{1}{n} \frac{m(m-1)}{2} \quad (\dagger) \approx -\frac{m^2}{2n} \end{aligned}$$

(*) We used $\ln(1 - \varepsilon) \approx -\varepsilon$ for $|\varepsilon| \ll 1$.

The Calculation.

A_i = no collision when i th ball is placed in a bin.

$$\Pr[A_i | A_{i-1} \cap \dots \cap A_1] = \left(1 - \frac{i-1}{n}\right).$$

no collision = $A_1 \cap \dots \cap A_m$.

Product rule:

$$\Pr[A_1 \cap \dots \cap A_m] = \Pr[A_1] \Pr[A_2 | A_1] \dots \Pr[A_m | A_1 \cap \dots \cap A_{m-1}]$$

$$\Rightarrow \Pr[\text{no collision}] = \left(1 - \frac{1}{n}\right) \dots \left(1 - \frac{m-1}{n}\right).$$

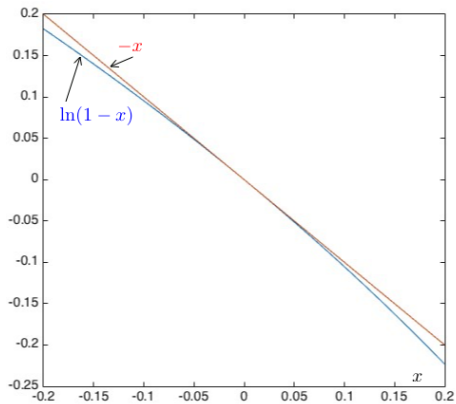
Hence,

$$\begin{aligned} \ln(\Pr[\text{no collision}]) &= \sum_{k=1}^{m-1} \ln\left(1 - \frac{k}{n}\right) \approx \sum_{k=1}^{m-1} \left(-\frac{k}{n}\right) (*) \\ &= -\frac{1}{n} \frac{m(m-1)}{2} (\dagger) \approx -\frac{m^2}{2n} \end{aligned}$$

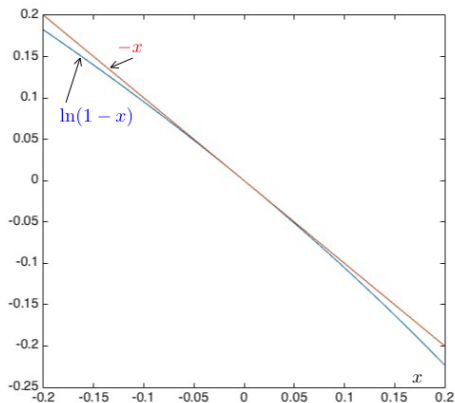
(*) We used $\ln(1 - \varepsilon) \approx -\varepsilon$ for $|\varepsilon| \ll 1$.

(†) $1 + 2 + \dots + m - 1 = (m - 1)m/2$.

Approximation

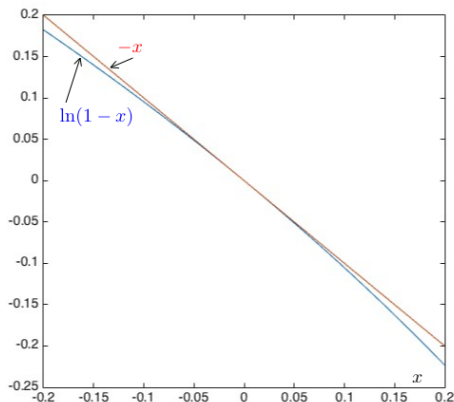


Approximation



$$\exp\{-x\} = 1 - x + \frac{1}{2!}x^2 + \dots \approx 1 - x, \text{ for } |x| \ll 1.$$

Approximation



$$\exp\{-x\} = 1 - x + \frac{1}{2!}x^2 + \dots \approx 1 - x, \text{ for } |x| \ll 1.$$

Hence, $-x \approx \ln(1-x)$ for $|x| \ll 1$.

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

$Pr[\text{collision}] \approx 1/2$ if $m \approx 1.2\sqrt{365} \approx 23$.

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

$Pr[\text{collision}] \approx 1/2$ if $m \approx 1.2\sqrt{365} \approx 23$.

If $m = 60$, we find that

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

$Pr[\text{collision}] \approx 1/2$ if $m \approx 1.2\sqrt{365} \approx 23$.

If $m = 60$, we find that

$$Pr[\text{no collision}] \approx \exp\left\{-\frac{m^2}{2n}\right\} = \exp\left\{-\frac{60^2}{2 \times 365}\right\} \approx 0.007.$$

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

$Pr[\text{collision}] \approx 1/2$ if $m \approx 1.2\sqrt{365} \approx 23$.

If $m = 60$, we find that

$$Pr[\text{no collision}] \approx \exp\left\{-\frac{m^2}{2n}\right\} = \exp\left\{-\frac{60^2}{2 \times 365}\right\} \approx 0.007.$$

If $m = 366$, then $Pr[\text{no collision}] =$

Today's your birthday, it's my birthday too..

Probability that m people all have different birthdays?

With $n = 365$, one finds

$Pr[\text{collision}] \approx 1/2$ if $m \approx 1.2\sqrt{365} \approx 23$.

If $m = 60$, we find that

$$Pr[\text{no collision}] \approx \exp\left\{-\frac{m^2}{2n}\right\} = \exp\left\{-\frac{60^2}{2 \times 365}\right\} \approx 0.007.$$

If $m = 366$, then $Pr[\text{no collision}] = 0$. (No approximation here!)

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Ball i in some bin, ball j chooses that bin with probability $1/n$.

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = \Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Ball i in some bin, ball j chooses that bin with probability $1/n$.

$$E[X] = \frac{m(m-1)}{2n} \approx \frac{m^2}{2n}.$$

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = \Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Ball i in some bin, ball j chooses that bin with probability $1/n$.

$$E[X] = \frac{m(m-1)}{2n} \approx \frac{m^2}{2n}.$$

For $m = \sqrt{n}$, $E[X] = 1/2$

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = \Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Ball i in some bin, ball j chooses that bin with probability $1/n$.

$$E[X] = \frac{m(m-1)}{2n} \approx \frac{m^2}{2n}.$$

For $m = \sqrt{n}$, $E[X] = 1/2$

Markov: $\Pr[X \geq c] \leq \frac{EX}{c}$.

Using linearity of expectation.

Experiment: m balls into n bins uniformly at random.

Random Variable:

X = Number of collisions between pairs of balls.

or number of pairs i and j where ball i and ball j are in same bin.

$$X_{ij} = 1 \{\text{balls } i, j \text{ in same bin}\}$$

$$X = \sum_{ij} X_{ij}$$

$$E[X_{ij}] = \Pr[\text{balls } i, j \text{ in same bin}] = \frac{1}{n}.$$

Ball i in some bin, ball j chooses that bin with probability $1/n$.

$$E[X] = \frac{m(m-1)}{2n} \approx \frac{m^2}{2n}.$$

$$\text{For } m = \sqrt{n}, E[X] = 1/2$$

$$\text{Markov: } \Pr[X \geq c] \leq \frac{EX}{c}.$$

$$\Pr[X \geq 1] \leq \frac{E[X]}{1} = 1/2.$$

Checksums!

Checksums!

Consider a set of m files.

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\}$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $\Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $\Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$.

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$Pr[\text{no collision}] \approx 1 - 10^{-3}$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$Pr[\text{no collision}] \approx 1 - 10^{-3} \Leftrightarrow m^2/(2n) \approx 10^{-3}$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$\begin{aligned} Pr[\text{no collision}] \approx 1 - 10^{-3} &\Leftrightarrow m^2/(2n) \approx 10^{-3} \\ &\Leftrightarrow 2n \approx m^2 10^3 \end{aligned}$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$\begin{aligned} Pr[\text{no collision}] \approx 1 - 10^{-3} &\Leftrightarrow m^2/(2n) \approx 10^{-3} \\ &\Leftrightarrow 2n \approx m^2 10^3 \Leftrightarrow 2^{b+1} \approx m^2 2^{10} \end{aligned}$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$Pr[\text{no collision}] \approx 1 - 10^{-3} \Leftrightarrow m^2/(2n) \approx 10^{-3}$$

$$\Leftrightarrow 2n \approx m^2 10^3 \Leftrightarrow 2^{b+1} \approx m^2 2^{10}$$

$$\Leftrightarrow b + 1 \approx 10 + 2 \log_2(m)$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$Pr[\text{no collision}] \approx 1 - 10^{-3} \Leftrightarrow m^2/(2n) \approx 10^{-3}$$

$$\Leftrightarrow 2n \approx m^2 10^3 \Leftrightarrow 2^{b+1} \approx m^2 2^{10}$$

$$\Leftrightarrow b+1 \approx 10 + 2 \log_2(m) \approx 10 + 2.9 \ln(m).$$

Checksums!

Consider a set of m files.

Each file has a checksum of b bits.

How large should b be for $Pr[\text{share a checksum}] \leq 10^{-3}$?

Claim: $b \geq 2.9 \ln(m) + 9$.

Proof:

Let $n = 2^b$ be the number of checksums.

We know $Pr[\text{no collision}] \approx \exp\{-m^2/(2n)\} \approx 1 - m^2/(2n)$. Hence,

$$\begin{aligned} Pr[\text{no collision}] \approx 1 - 10^{-3} &\Leftrightarrow m^2/(2n) \approx 10^{-3} \\ &\Leftrightarrow 2n \approx m^2 10^3 \Leftrightarrow 2^{b+1} \approx m^2 2^{10} \\ &\Leftrightarrow b+1 \approx 10 + 2 \log_2(m) \approx 10 + 2.9 \ln(m). \end{aligned}$$

Note: $\log_2(x) = \log_2(e) \ln(x) \approx 1.44 \ln(x)$.

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.



Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.



Theorem:

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.



Theorem: If you buy m boxes,

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.



Theorem: If you buy m boxes,

(a) $Pr[\text{miss one specific item}] \approx e^{-\frac{m}{n}}$

Coupon Collector Problem.

There are n different baseball cards.

(Brian Wilson, Jackie Robinson, Roger Hornsby, ...)

One random baseball card in each cereal box.



Theorem: If you buy m boxes,

(a) $Pr[\text{miss one specific item}] \approx e^{-\frac{m}{n}}$

(b) $Pr[\text{miss any one of the items}] \leq ne^{-\frac{m}{n}}$.

Coupon Collector Problem: Analysis.

Event $A_m =$ 'fail to get Brian Wilson in m cereal boxes'

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ...

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$Pr[A_m] = (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n})$$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$\begin{aligned}Pr[A_m] &= (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n}) \\ &= (1 - \frac{1}{n})^m\end{aligned}$$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$\begin{aligned}Pr[A_m] &= (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n}) \\&= (1 - \frac{1}{n})^m \\ \ln(Pr[A_m]) &= m \ln(1 - \frac{1}{n}) \approx\end{aligned}$$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$\begin{aligned}Pr[A_m] &= (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n}) \\&= (1 - \frac{1}{n})^m \\ \ln(Pr[A_m]) &= m \ln(1 - \frac{1}{n}) \approx m \times (-\frac{1}{n})\end{aligned}$$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$\begin{aligned}Pr[A_m] &= (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n}) \\ &= (1 - \frac{1}{n})^m\end{aligned}$$

$$\ln(Pr[A_m]) = m \ln(1 - \frac{1}{n}) \approx m \times (-\frac{1}{n})$$

$$Pr[A_m] \approx \exp\{-\frac{m}{n}\}.$$

Coupon Collector Problem: Analysis.

Event A_m = 'fail to get Brian Wilson in m cereal boxes'

Fail the first time: $(1 - \frac{1}{n})$

Fail the second time: $(1 - \frac{1}{n})$

And so on ... for m times. Hence,

$$\begin{aligned}Pr[A_m] &= (1 - \frac{1}{n}) \times \dots \times (1 - \frac{1}{n}) \\ &= (1 - \frac{1}{n})^m\end{aligned}$$

$$\ln(Pr[A_m]) = m \ln(1 - \frac{1}{n}) \approx m \times (-\frac{1}{n})$$

$$Pr[A_m] \approx \exp\{-\frac{m}{n}\}.$$

For $p_m = \frac{1}{2}$, we need around $n \ln 2 \approx 0.69n$ boxes.

Collect all cards?

Experiment: Choose m cards at random with replacement.

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ' , for $k = 1, \dots, n$

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Probability of failing to get at least one of these n players:

$$p := Pr[E_1 \cup E_2 \cdots \cup E_n]$$

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Probability of failing to get at least one of these n players:

$$p := Pr[E_1 \cup E_2 \cdots \cup E_n]$$

How does one estimate p ?

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Probability of failing to get at least one of these n players:

$$p := \Pr[E_1 \cup E_2 \cdots \cup E_n]$$

How does one estimate p ? **Union Bound:**

$$p = \Pr[E_1 \cup E_2 \cdots \cup E_n] \leq \Pr[E_1] + \Pr[E_2] \cdots \Pr[E_n].$$

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Probability of failing to get at least one of these n players:

$$p := \Pr[E_1 \cup E_2 \cdots \cup E_n]$$

How does one estimate p ? **Union Bound:**

$$p = \Pr[E_1 \cup E_2 \cdots \cup E_n] \leq \Pr[E_1] + \Pr[E_2] \cdots \Pr[E_n].$$

$$\Pr[E_k] \approx e^{-\frac{m}{n}}, k = 1, \dots, n.$$

Collect all cards?

Experiment: Choose m cards at random with replacement.

Events: $E_k =$ 'fail to get player k ', for $k = 1, \dots, n$

Probability of failing to get at least one of these n players:

$$p := \Pr[E_1 \cup E_2 \cdots \cup E_n]$$

How does one estimate p ? **Union Bound:**

$$p = \Pr[E_1 \cup E_2 \cdots \cup E_n] \leq \Pr[E_1] + \Pr[E_2] \cdots \Pr[E_n].$$

$$\Pr[E_k] \approx e^{-\frac{m}{n}}, k = 1, \dots, n.$$

Plug in and get

$$p \leq ne^{-\frac{m}{n}}.$$

Collect all cards?

Thus,

$$Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Collect all cards?

Thus,

$$Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Hence,

$$Pr[\text{missing at least one card}] \leq p \text{ when } m \geq n \ln\left(\frac{n}{p}\right).$$

Collect all cards?

Thus,

$$\Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Hence,

$$\Pr[\text{missing at least one card}] \leq p \text{ when } m \geq n \ln\left(\frac{n}{p}\right).$$

To get $p = 1/2$, set $m = n \ln(2n)$.

Collect all cards?

Thus,

$$\Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Hence,

$$\Pr[\text{missing at least one card}] \leq p \text{ when } m \geq n \ln\left(\frac{n}{p}\right).$$

To get $p = 1/2$, set $m = n \ln(2n)$.

$$(p \leq ne^{-\frac{m}{n}} \leq ne^{-\ln(n/p)} \leq n\left(\frac{p}{n}\right) \leq p.)$$

Collect all cards?

Thus,

$$\Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Hence,

$$\Pr[\text{missing at least one card}] \leq p \text{ when } m \geq n \ln\left(\frac{n}{p}\right).$$

To get $p = 1/2$, set $m = n \ln(2n)$.

$$(p \leq ne^{-\frac{m}{n}} \leq ne^{-\ln(n/p)} \leq n\left(\frac{p}{n}\right) \leq p.)$$

E.g., $n = 10^2 \Rightarrow m = 530$;

Collect all cards?

Thus,

$$\Pr[\text{missing at least one card}] \leq ne^{-\frac{m}{n}}.$$

Hence,

$$\Pr[\text{missing at least one card}] \leq p \text{ when } m \geq n \ln\left(\frac{n}{p}\right).$$

To get $p = 1/2$, set $m = n \ln(2n)$.

$$(p \leq ne^{-\frac{m}{n}} \leq ne^{-\ln(n/p)} \leq n\left(\frac{p}{n}\right) \leq p.)$$

E.g., $n = 10^2 \Rightarrow m = 530$; $n = 10^3 \Rightarrow m = 7600$.

Time to collect coupons

X -time to get n coupons.

Time to collect coupons

X - time to get n coupons.

X_1 - time to get first coupon.

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$.

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk"}]$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$?

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric !

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? **Geometric ! !**

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? **Geometric !!!**

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric !!! $\implies E[X_2] = \frac{1}{p} =$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric !!! $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}}$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric !!! $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}$.

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i]$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p}$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1},$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, \dots, n.$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, \dots, n.$$

$$E[X] = E[X_1] + \dots + E[X_n] =$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$

$E[X_2]$? Geometric !!! $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}$.

$Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$

$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, \dots, n$.

$$E[X] = E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1}$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, \dots, n.$$

$$\begin{aligned} E[X] &= E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1} \\ &= n\left(1 + \frac{1}{2} + \dots + \frac{1}{n}\right) =: nH(n) \end{aligned}$$

Time to collect coupons

X -time to get n coupons.

X_1 - time to get first coupon. Note: $X_1 = 1$. $E(X_1) = 1$.

X_2 - time to get second coupon after getting first.

$$\Pr[\text{"get second coupon"} | \text{"got milk first coupon"}] = \frac{n-1}{n}$$

$$E[X_2]? \text{ Geometric !!! } \implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}.$$

$$\Pr[\text{"getting } i\text{th coupon"} | \text{"got } i-1 \text{rst coupons"}] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, \dots, n.$$

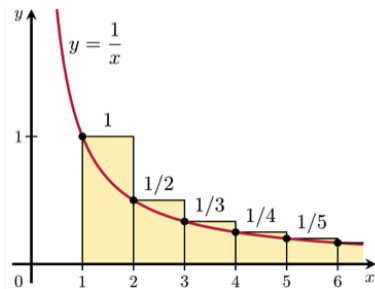
$$\begin{aligned} E[X] &= E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1} \\ &= n \left(1 + \frac{1}{2} + \dots + \frac{1}{n} \right) =: nH(n) \approx n(\ln n + \gamma) \end{aligned}$$

Review: Harmonic sum

$$H(n) = 1 + \frac{1}{2} + \cdots + \frac{1}{n} \approx \int_1^n \frac{1}{x} dx = \ln(n).$$

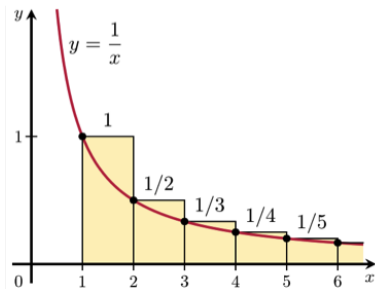
Review: Harmonic sum

$$H(n) = 1 + \frac{1}{2} + \cdots + \frac{1}{n} \approx \int_1^n \frac{1}{x} dx = \ln(n).$$



Review: Harmonic sum

$$H(n) = 1 + \frac{1}{2} + \cdots + \frac{1}{n} \approx \int_1^n \frac{1}{x} dx = \ln(n).$$



A good approximation is

$$H(n) \approx \ln(n) + \gamma \text{ where } \gamma \approx 0.58 \text{ (Euler-Mascheroni constant).}$$

Simplest..

Load balance: m balls in n bins.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin:

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized!

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random?

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n .

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n . Uh Oh!

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n . Uh Oh!

Max load with probability $\geq 1 - \delta$?

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n . Uh Oh!

Max load with probability $\geq 1 - \delta$?

$\delta = \frac{1}{n^c}$ for today.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n . Uh Oh!

Max load with probability $\geq 1 - \delta$?

$\delta = \frac{1}{n^c}$ for today. c is 1 or 2.

Simplest..

Load balance: m balls in n bins.

For simplicity: n balls in n bins.

Round robin: load 1 !

Centralized! Not so good.

Uniformly at random? Average load 1.

Max load?

n . Uh Oh!

Max load with probability $\geq 1 - \delta$?

$\delta = \frac{1}{n^c}$ for today. c is 1 or 2.

Balls in bins.

For each of n balls, choose random bin:

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k \quad \text{and} \quad \binom{n}{k} \text{ subsets } S.$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k$ and $\binom{n}{k}$ subsets S .

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k \quad \text{and} \quad \binom{n}{k} \text{ subsets } S.$$

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Choose k , so that $\Pr[X_i \geq k] \leq \frac{1}{n^2}$.

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k \quad \text{and} \quad \binom{n}{k} \text{ subsets } S.$$

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Choose k , so that $\Pr[X_i \geq k] \leq \frac{1}{n^2}$.

$$\Pr[\text{any } X_i \geq k] \leq n \times \frac{1}{n^2}$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k \quad \text{and} \quad \binom{n}{k} \text{ subsets } S.$$

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Choose k , so that $\Pr[X_i \geq k] \leq \frac{1}{n^2}$.

$$\Pr[\text{any } X_i \geq k] \leq n \times \frac{1}{n^2} = \frac{1}{n}$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k$ and $\binom{n}{k}$ subsets S .

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Choose k , so that $\Pr[X_i \geq k] \leq \frac{1}{n^2}$.

$$\Pr[\text{any } X_i \geq k] \leq n \times \frac{1}{n^2} = \frac{1}{n} \rightarrow \text{max load} \leq k \text{ w.p. } \geq 1 - \frac{1}{n}$$

Balls in bins.

For each of n balls, choose random bin: X_i balls in bin i .

$$\Pr[X_i \geq k] \leq \sum_{S \subseteq [n], |S|=k} \Pr[\text{balls in } S \text{ chooses bin } i]$$

From Union Bound: $\Pr[\cup_i A_i] \leq \sum_i \Pr[A_i]$

$\Pr[\text{balls in } S \text{ chooses bin } i] = \left(\frac{1}{n}\right)^k$ and $\binom{n}{k}$ subsets S .

$$\begin{aligned} \Pr[X_i \geq k] &\leq \binom{n}{k} \left(\frac{1}{n}\right)^k \\ &\leq \frac{n^k}{k!} \left(\frac{1}{n}\right)^k = \frac{1}{k!} \end{aligned}$$

Choose k , so that $\Pr[X_i \geq k] \leq \frac{1}{n^2}$.

$$\Pr[\text{any } X_i \geq k] \leq n \times \frac{1}{n^2} = \frac{1}{n} \rightarrow \text{max load} \leq k \text{ w.p. } \geq 1 - \frac{1}{n}$$

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Also: $k = \Theta(\log n / \log \log n)$ suffices as well.

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Also: $k = \Theta(\log n / \log \log n)$ suffices as well.

$$k^k \rightarrow n^c.$$

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Also: $k = \Theta(\log n / \log \log n)$ suffices as well.

$$k^k \rightarrow n^c.$$

Actually Max load is $\Theta(\log n / \log \log n)$ w.h.p.

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Also: $k = \Theta(\log n / \log \log n)$ suffices as well.

$$k^k \rightarrow n^c.$$

Actually Max load is $\Theta(\log n / \log \log n)$ w.h.p.

(W.h.p. - means with probability at least $1 - O(1/n^c)$ for today.)

Solving for k

$$\Pr[X_i \geq k] \leq \frac{1}{k!} \leq 1/n^2?$$

What is upper bound on max-load k ?

Lemma: Max load is $\Theta(\log n)$ with probability $\geq 1 - \frac{1}{n}$.

$$k! \geq n^2 \text{ for } k = 2e \log n$$

(Recall $k! \geq (\frac{k}{e})^k$.)

$$\implies \frac{1}{k!} \leq \left(\frac{e}{k}\right)^k \leq \left(\frac{1}{2 \log n}\right)^k$$

If $\log n \geq 1$, then $k = 2e \log n$ suffices.

Also: $k = \Theta(\log n / \log \log n)$ suffices as well.

$$k^k \rightarrow n^c.$$

Actually Max load is $\Theta(\log n / \log \log n)$ w.h.p.

(W.h.p. - means with probability at least $1 - O(1/n^c)$ for today.)

Better than variance based methods...

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$\Pr[X > n] = \Pr[\text{first } n \text{ flips are } T] =$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$\Pr[X > n] = \Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$Pr[X > n] = Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$Pr[X > n] = Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

Proof:

$$Pr[X > n + m | X > n] =$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$\Pr[X > n] = \Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$\Pr[X > n + m | X > n] = \Pr[X > m], m, n \geq 0.$$

Proof:

$$\Pr[X > n + m | X > n] = \frac{\Pr[X > n + m \text{ and } X > n]}{\Pr[X > n]}$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$\Pr[X > n] = \Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$\Pr[X > n + m | X > n] = \Pr[X > m], m, n \geq 0.$$

Proof:

$$\begin{aligned} \Pr[X > n + m | X > n] &= \frac{\Pr[X > n + m \text{ and } X > n]}{\Pr[X > n]} \\ &= \frac{\Pr[X > n + m]}{\Pr[X > n]} \end{aligned}$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$Pr[X > n] = Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

Proof:

$$\begin{aligned} Pr[X > n + m | X > n] &= \frac{Pr[X > n + m \text{ and } X > n]}{Pr[X > n]} \\ &= \frac{Pr[X > n + m]}{Pr[X > n]} \\ &= \frac{(1 - p)^{n+m}}{(1 - p)^n} = \end{aligned}$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$Pr[X > n] = Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

Proof:

$$\begin{aligned} Pr[X > n + m | X > n] &= \frac{Pr[X > n + m \text{ and } X > n]}{Pr[X > n]} \\ &= \frac{Pr[X > n + m]}{Pr[X > n]} \\ &= \frac{(1 - p)^{n+m}}{(1 - p)^n} = (1 - p)^m \end{aligned}$$

Geometric Distribution: Memoryless

Let X be $G(p)$. Then, for $n \geq 0$,

$$Pr[X > n] = Pr[\text{first } n \text{ flips are } T] = (1 - p)^n.$$

Theorem

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

Proof:

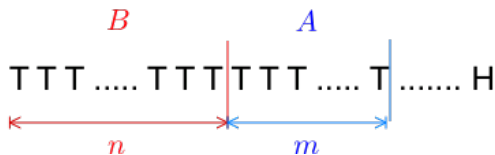
$$\begin{aligned} Pr[X > n + m | X > n] &= \frac{Pr[X > n + m \text{ and } X > n]}{Pr[X > n]} \\ &= \frac{Pr[X > n + m]}{Pr[X > n]} \\ &= \frac{(1 - p)^{n+m}}{(1 - p)^n} = (1 - p)^m \\ &= Pr[X > m]. \end{aligned}$$

Geometric Distribution: Memoryless - Interpretation

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$

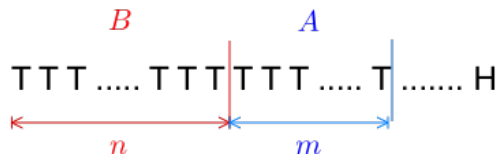
Geometric Distribution: Memoryless - Interpretation

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$



Geometric Distribution: Memoryless - Interpretation

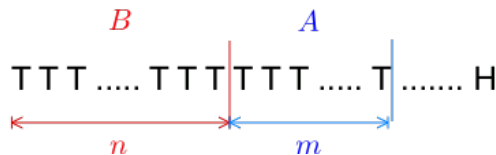
$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$



$$Pr[X > n + m | X > n] = Pr[A|B] = Pr[A] = Pr[X > m].$$

Geometric Distribution: Memoryless - Interpretation

$$Pr[X > n + m | X > n] = Pr[X > m], m, n \geq 0.$$



$$Pr[X > n + m | X > n] = Pr[A|B] = Pr[A] = Pr[X > m].$$

The coin is memoryless, therefore, so is X .

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Proof: One has

$$E[X] = \sum_{i=1}^{\infty} i \times \Pr[X = i]$$

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Proof: One has

$$\begin{aligned} E[X] &= \sum_{i=1}^{\infty} i \times \Pr[X = i] \\ &= \sum_{i=1}^{\infty} i \{ \Pr[X \geq i] - \Pr[X \geq i+1] \} \end{aligned}$$

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Proof: One has

$$\begin{aligned} E[X] &= \sum_{i=1}^{\infty} i \times \Pr[X = i] \\ &= \sum_{i=1}^{\infty} i \{ \Pr[X \geq i] - \Pr[X \geq i+1] \} \\ &= \sum_{i=1}^{\infty} \{ i \times \Pr[X \geq i] - i \times \Pr[X \geq i+1] \} \end{aligned}$$

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Proof: One has

$$\begin{aligned} E[X] &= \sum_{i=1}^{\infty} i \times \Pr[X = i] \\ &= \sum_{i=1}^{\infty} i \{ \Pr[X \geq i] - \Pr[X \geq i+1] \} \\ &= \sum_{i=1}^{\infty} \{ i \times \Pr[X \geq i] - i \times \Pr[X \geq i+1] \} \\ &= \sum_{i=1}^{\infty} \{ i \times \Pr[X \geq i] - (i-1) \times \Pr[X \geq i] \} \end{aligned}$$

Expected Value of Integer RV

Theorem: For a r.v. X that takes values in $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

Proof: One has

$$\begin{aligned} E[X] &= \sum_{i=1}^{\infty} i \times \Pr[X = i] \\ &= \sum_{i=1}^{\infty} i \{ \Pr[X \geq i] - \Pr[X \geq i+1] \} \\ &= \sum_{i=1}^{\infty} \{ i \times \Pr[X \geq i] - i \times \Pr[X \geq i+1] \} \\ &= \sum_{i=1}^{\infty} \{ i \times \Pr[X \geq i] - (i-1) \times \Pr[X \geq i] \} \\ &= \sum_{i=1}^{\infty} \Pr[X \geq i]. \end{aligned}$$



Geometric Distribution: Yet another look

Theorem: For a r.v. X that takes the values $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} Pr[X \geq i].$$

Geometric Distribution: Yet another look

Theorem: For a r.v. X that takes the values $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

If $X = G(p)$, then $\Pr[X \geq i] = \Pr[X > i - 1] = (1 - p)^{i-1}$.

Geometric Distribution: Yet another look

Theorem: For a r.v. X that takes the values $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

If $X = G(p)$, then $\Pr[X \geq i] = \Pr[X > i - 1] = (1 - p)^{i-1}$.

Hence,

$$E[X] = \sum_{i=1}^{\infty} (1 - p)^{i-1} = \sum_{i=0}^{\infty} (1 - p)^i$$

Geometric Distribution: Yet another look

Theorem: For a r.v. X that takes the values $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

If $X = G(p)$, then $\Pr[X \geq i] = \Pr[X > i - 1] = (1 - p)^{i-1}$.

Hence,

$$E[X] = \sum_{i=1}^{\infty} (1 - p)^{i-1} = \sum_{i=0}^{\infty} (1 - p)^i = \frac{1}{1 - (1 - p)} =$$

Geometric Distribution: Yet another look

Theorem: For a r.v. X that takes the values $\{0, 1, 2, \dots\}$, one has

$$E[X] = \sum_{i=1}^{\infty} \Pr[X \geq i].$$

If $X = G(p)$, then $\Pr[X \geq i] = \Pr[X > i - 1] = (1 - p)^{i-1}$.

Hence,

$$E[X] = \sum_{i=1}^{\infty} (1 - p)^{i-1} = \sum_{i=0}^{\infty} (1 - p)^i = \frac{1}{1 - (1 - p)} = \frac{1}{p}.$$

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson?

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter?

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Now:

arrival for X happens with probability λ/n

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Now:

arrival for X happens with probability λ/n

arrival for Y happens with probability μ/n

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Now:

arrival for X happens with probability λ/n

arrival for Y happens with probability μ/n

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Now:

arrival for X happens with probability λ/n

arrival for Y happens with probability μ/n

So, we get limit $n \rightarrow \infty$ is $B(n, (\lambda + \mu)/n)$.

Sum of Poisson Random Variables.

For $X = P(\lambda)$ and $Y = P(\mu)$, what is $X + Y$?

Poisson? Yes.

What parameter? $\lambda + \mu$.

Why?

$P(\lambda)$ is limit $n \rightarrow \infty$ of $B(n, \lambda/n)$.

Recall Derivation:

break interval into n intervals

and each has arrival with probability λ/n .

Now:

arrival for X happens with probability λ/n

arrival for Y happens with probability μ/n

So, we get limit $n \rightarrow \infty$ is $B(n, (\lambda + \mu)/n)$.

Details: both could arrive with probability $\lambda\mu/n^2$.

But this goes to zero as $n \rightarrow \infty$.

(Like λ^2/n^2 in previous derivation)

Linear Regression: Preamble

Linear Regression: Preamble

The “best” guess about Y ,

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$E[(Y - a)^2] = E[(Y - E[Y] + E[Y] - a)^2]$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \geq E[\hat{Y}^2]. \end{aligned}$$

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \geq E[\hat{Y}^2]. \end{aligned}$$

Hence, $E[(Y - a)^2] \geq E[(Y - E[Y])^2], \forall a$.

Linear Regression: Preamble

The “best” guess about Y , if we know only the distribution of Y , is $E[Y]$.

If “best” is Mean Squared Error.

More precisely, the value of a that minimizes $E[(Y - a)^2]$ is $a = E[Y]$.

Proof:

Let $\hat{Y} := Y - E[Y]$.

Then, $E[\hat{Y}] = E[Y - E[Y]] = E[Y] - E[Y] = 0$.

So, $E[\hat{Y}c] = 0, \forall c$. Now,

$$\begin{aligned} E[(Y - a)^2] &= E[(Y - E[Y] + E[Y] - a)^2] \\ &= E[(\hat{Y} + c)^2] \text{ with } c = E[Y] - a \\ &= E[\hat{Y}^2 + 2\hat{Y}c + c^2] = E[\hat{Y}^2] + 2E[\hat{Y}c] + c^2 \\ &= E[\hat{Y}^2] + 0 + c^2 \geq E[\hat{Y}^2]. \end{aligned}$$

Hence, $E[(Y - a)^2] \geq E[(Y - E[Y])^2], \forall a$.



Linear Regression: Preamble

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

The simplest function $g(X)$ is a constant that does not depend of X .

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

The simplest function $g(X)$ is a constant that does not depend of X .

The next simplest function is linear: $g(X) = a + bX$.

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

The simplest function $g(X)$ is a constant that does not depend of X .

The next simplest function is linear: $g(X) = a + bX$.

What is the best linear function?

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

The simplest function $g(X)$ is a constant that does not depend of X .

The next simplest function is linear: $g(X) = a + bX$.

What is the best linear function? That is our next topic.

Linear Regression: Preamble

Thus, if we want to guess the value of Y , we choose $E[Y]$.

Now assume we make some observation X related to Y .

How do we use that observation to improve our guess about Y ?

The idea is to use a function $g(X)$ of the observation to estimate Y .

The simplest function $g(X)$ is a constant that does not depend of X .

The next simplest function is linear: $g(X) = a + bX$.

What is the best linear function? That is our next topic.

A bit later, we will consider a general function $g(X)$.

Linear Regression: Motivation

Linear Regression: Motivation

Example 1: 100 people.

Linear Regression: Motivation

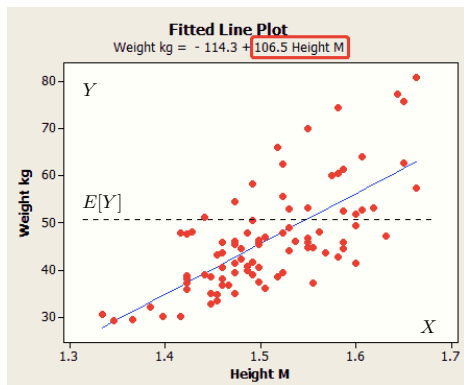
Example 1: 100 people.

Let $(X_n, Y_n) = (\text{height}, \text{weight})$ of person n , for $n = 1, \dots, 100$:

Linear Regression: Motivation

Example 1: 100 people.

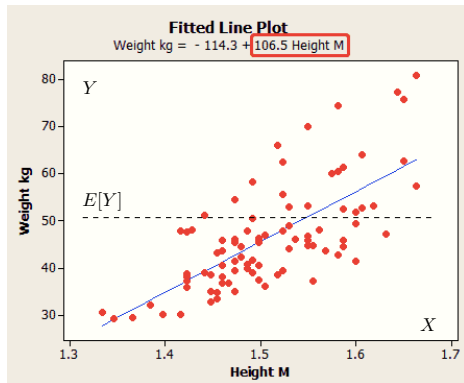
Let $(X_n, Y_n) = (\text{height}, \text{weight})$ of person n , for $n = 1, \dots, 100$:



Linear Regression: Motivation

Example 1: 100 people.

Let $(X_n, Y_n) = (\text{height}, \text{weight})$ of person n , for $n = 1, \dots, 100$:

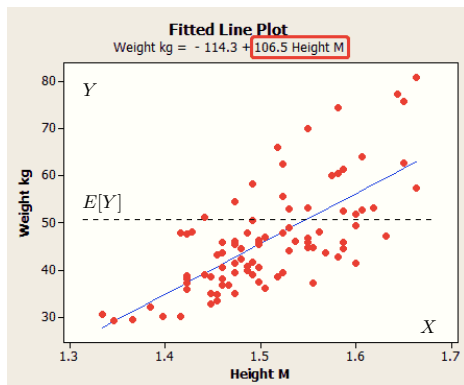


The blue line is $Y = -114.3 + 106.5X$. (X in meters, Y in kg.)

Linear Regression: Motivation

Example 1: 100 people.

Let $(X_n, Y_n) = (\text{height}, \text{weight})$ of person n , for $n = 1, \dots, 100$:



The blue line is $Y = -114.3 + 106.5X$. (X in meters, Y in kg.)

Best linear fit: [Linear Regression](#).

Motivation

Example 2: 15 people.

Motivation

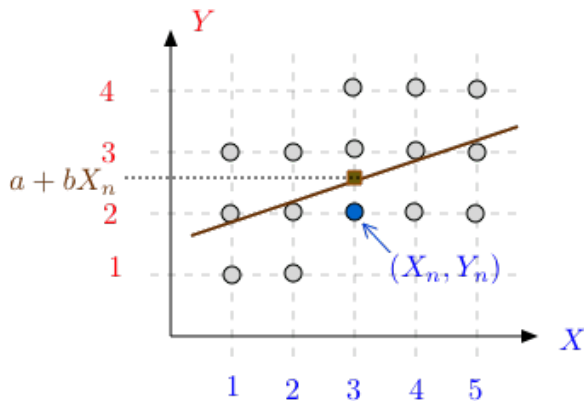
Example 2: 15 people.

We look at two attributes: (X_n, Y_n) of person n , for $n = 1, \dots, 15$:

Motivation

Example 2: 15 people.

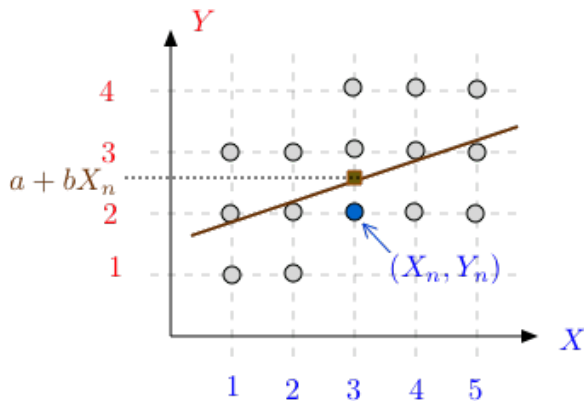
We look at two attributes: (X_n, Y_n) of person n , for $n = 1, \dots, 15$:



Motivation

Example 2: 15 people.

We look at two attributes: (X_n, Y_n) of person n , for $n = 1, \dots, 15$:



The line $Y = a + bX$ is the linear regression.

LLSE

$LLSE[Y|X]$ - best guess for Y given X .

LLSE

$LLSE[Y|X]$ - best guess for Y given X .

Theorem

LLSE

$LLSE[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

LLSE

$LLSE[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

LLSE

LLSE $[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

LLSE

LLSE[$Y|X$] - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

LLSE

LLSE $[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$,

LLSE

LLSE $[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

LLSE

LLSE[$Y|X$] - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β ,

LLSE

LLSE[$Y|X$] - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

LLSE

LLSE $[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$.

LLSE

LLSE[Y|X] - best guess for Y given X.

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - a - bX = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - a - bX)] = 0, \forall a, b$.

LLSE

LLSE[Y|X] - best guess for Y given X.

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$E[(Y - \alpha - \beta X)^2] = E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2]$$

LLSE

LLSE[Y|X] - best guess for Y given X.

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \end{aligned}$$

LLSE

LLSE[$Y|X$] - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]). \quad E[Y - \hat{Y}] = 0 \text{ by linearity.}$$

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

LLSE

LLSE[Y|X] - best guess for Y given X.

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$, for all (a, b) .

LLSE

LLSE $[Y|X]$ - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - \alpha - \beta X)^2]$, for all (α, β) .

Thus \hat{Y} is the LLSE.

LLSE

LLSE[$Y|X$] - best guess for Y given X .

Theorem

Consider two RVs X, Y with a given distribution $Pr[X = x, Y = y]$.

Then,

$$L[Y|X] = \hat{Y} = E[Y] + \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X]).$$

Proof 1:

$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}(X)}(X - E[X])$. $E[Y - \hat{Y}] = 0$ by linearity.

Also, $E[(Y - \hat{Y})X] = 0$, after a bit of algebra. (next slide)

Combine brown inequalities: $E[(Y - \hat{Y})(c + dX)] = 0$ for any c, d .

Since: $\hat{Y} = \alpha + \beta X$ for some α, β , so $\exists c, d$ s.t. $\hat{Y} - \alpha - \beta X = c + dX$.

Then, $E[(Y - \hat{Y})(\hat{Y} - \alpha - \beta X)] = 0, \forall \alpha, \beta$. Now,

$$\begin{aligned} E[(Y - \alpha - \beta X)^2] &= E[(Y - \hat{Y} + \hat{Y} - \alpha - \beta X)^2] \\ &= E[(Y - \hat{Y})^2] + E[(\hat{Y} - \alpha - \beta X)^2] + 0 \geq E[(Y - \hat{Y})^2]. \end{aligned}$$

This shows that $E[(Y - \hat{Y})^2] \leq E[(Y - a - bX)^2]$, for all (a, b) .

Thus \hat{Y} is the LLSE. □

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]} (X - E[X]).$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

Now,

$$\begin{aligned} & E[(Y - \hat{Y})(X - E[X])] \\ &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}[X]} E[(X - E[X])(X - E[X])] \end{aligned}$$

A Bit of Algebra

$$Y - \hat{Y} = (Y - E[Y]) - \frac{\text{cov}(X, Y)}{\text{var}[X]}(X - E[X]).$$

Hence, $E[Y - \hat{Y}] = 0$. We want to show that $E[(Y - \hat{Y})X] = 0$.

Note that

$$E[(Y - \hat{Y})X] = E[(Y - \hat{Y})(X - E[X])],$$

because $E[(Y - \hat{Y})E[X]] = 0$.

Now,

$$\begin{aligned} E[(Y - \hat{Y})(X - E[X])] &= E[(Y - E[Y])(X - E[X])] - \frac{\text{cov}(X, Y)}{\text{var}[X]} E[(X - E[X])(X - E[X])] \\ &=^{(*)} \text{cov}(X, Y) - \frac{\text{cov}(X, Y)}{\text{var}[X]} \text{var}[X] = 0. \quad \square \end{aligned}$$

(*) Recall that $\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$ and $\text{var}[X] = E[(X - E[X])^2]$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Simple Total Probability: $Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B]$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Simple Total Probability: $Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Simple Total Probability: $Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$.

Simple Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$.

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Simple Total Probability: $Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$.

Simple Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$.

Bayes Rule: $Pr[A|B] = \frac{Pr[B|A]Pr[A]}{Pr[B]}$

Discrete Probability.

Probability Space: Ω , $Pr : \Omega \rightarrow [0, 1]$, $\sum_{\omega \in \Omega} Pr(\omega) = 1$.

Events: $A \subset \Omega$.

Simple Total Probability: $Pr[B] = Pr[A \cap B] + Pr[\bar{A} \cap B]$.

Conditional Probability: $Pr[A|B] = \frac{Pr[A \cap B]}{Pr[B]}$.

Simple Product Rule: $Pr[A \cap B] = Pr[A|B]Pr[B]$.

Bayes Rule: $Pr[A|B] = \frac{Pr[B|A]Pr[A]}{Pr[B]}$

Inference:

Have one of two coins. Flip coin, which coin do you have?

Got positive test result. What is probability you have disease?

Random Variables

Random Variables: $X : \Omega \rightarrow \mathcal{R}$.

Random Variables

Random Variables: $X : \Omega \rightarrow \mathcal{R}$.

Random Variables

Random Variables: $X : \Omega \rightarrow \mathcal{R}$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

Random Variables

Random Variables: $X : \Omega \rightarrow \mathcal{R}$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a aPr[X = a] = \sum_{\omega \in \Omega} X(\omega)Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

$E[X] = \frac{n+1}{2}$, $Var(X) = \frac{n^2-1}{12}$.

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

$E[X] = \frac{n+1}{2}$, $Var(X) = \frac{n^2-1}{12}$.

Geometric: $X \sim G(p)$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

$E[X] = \frac{n+1}{2}$, $Var(X) = \frac{n^2-1}{12}$.

Geometric: $X \sim G(p)$ $Pr[X = i] = (1-p)^{i-1} p$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

$E[X] = \frac{n+1}{2}$, $Var(X) = \frac{n^2-1}{12}$.

Geometric: $X \sim G(p)$ $Pr[X = i] = (1-p)^{i-1} p$

$E(X) = \frac{1}{p}$, $Var(X) = \frac{1-p}{p^2}$

Random Variables

Random Variables: $X : \Omega \rightarrow R$.

Distribution: $Pr[X = a] = \sum_{\omega: X(\omega)=a} Pr(\omega)$

X and Y independent \iff all associated events are independent.

Expectation: $E[X] = \sum_a a Pr[X = a] = \sum_{\omega \in \Omega} X(\omega) Pr(\omega)$.

Linearity: $E[X + Y] = E[X] + E[Y]$.

Variance: $Var(X) = E[(X - E[X])^2] = E[X^2] - (E(X))^2$

For independent X, Y , $Var(X + Y) = Var(X) + Var(Y)$.

Also: $Var(cX) = c^2 Var(X)$ and $Var(X + b) = Var(X)$.

Poisson: $X \sim P(\lambda)$ $Pr[X = i] = e^{-\lambda} \frac{\lambda^i}{i!}$.

$E(X) = \lambda$, $Var(X) = \lambda$.

Binomial: $X \sim B(n, p)$ $Pr[X = i] = \binom{n}{i} p^i (1-p)^{n-i}$

$E(X) = np$, $Var(X) = np(1-p)$

Uniform: $X \sim U\{1, \dots, n\}$ $\forall i \in [1, n], Pr[X = i] = \frac{1}{n}$.

$E[X] = \frac{n+1}{2}$, $Var(X) = \frac{n^2-1}{12}$.

Geometric: $X \sim G(p)$ $Pr[X = i] = (1-p)^{i-1} p$

$E(X) = \frac{1}{p}$, $Var(X) = \frac{1-p}{p^2}$

Note: Probability Mass Function \equiv Distribution.