CS70.

- 1. Random Variables: Brief Review
- 2. Joint Distributions.
- 3. Linearity of Expectation

# Random Variables: Definitions Definition

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X:\Omega\to\Re$ .

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

(b) For  $A \subset \Re$ , one defines

$$X^{-1}(A) := \{\omega \in \Omega \mid X(\omega) \in A\}.$$

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

(b) For  $A \subset \Re$ , one defines

$$X^{-1}(A) := \{ \omega \in \Omega \mid X(\omega) \in A \}.$$

(c) The probability that X = a is defined as

$$Pr[X = a] = Pr[X^{-1}(a)].$$

#### **Definition**

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X:\Omega\to\Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

(b) For  $A \subset \Re$ , one defines

$$X^{-1}(A) := \{ \omega \in \Omega \mid X(\omega) \in A \}.$$

(c) The probability that X = a is defined as

$$Pr[X = a] = Pr[X^{-1}(a)].$$

(d) The probability that  $X \in A$  is defined as

$$Pr[X \in A] = Pr[X^{-1}(A)].$$

#### Definition

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

(b) For  $A \subset \mathfrak{R}$ , one defines

$$X^{-1}(A) := \{ \omega \in \Omega \mid X(\omega) \in A \}.$$

(c) The probability that X = a is defined as

$$Pr[X = a] = Pr[X^{-1}(a)].$$

(d) The probability that  $X \in A$  is defined as

$$Pr[X \in A] = Pr[X^{-1}(A)].$$

(e) The distribution of a random variable X, is

$$\{(a, Pr[X = a]) : a \in \mathscr{A}\},$$

where  $\mathscr{A}$  is the *range* of X.

#### Definition

A random variable, X, for a random experiment with sample space  $\Omega$  is a function  $X : \Omega \to \Re$ .

Thus,  $X(\cdot)$  assigns a real number  $X(\omega)$  to each  $\omega \in \Omega$ .

#### **Definitions**

(a) For  $a \in \Re$ , one defines

$$X^{-1}(a) := \{ \omega \in \Omega \mid X(\omega) = a \}.$$

(b) For  $A \subset \mathfrak{R}$ , one defines

$$X^{-1}(A) := \{ \omega \in \Omega \mid X(\omega) \in A \}.$$

(c) The probability that X = a is defined as

$$Pr[X = a] = Pr[X^{-1}(a)].$$

(d) The probability that  $X \in A$  is defined as

$$Pr[X \in A] = Pr[X^{-1}(A)].$$

(e) The distribution of a random variable X, is

$$\{(a, Pr[X = a]) : a \in \mathscr{A}\},$$

where  $\mathscr{A}$  is the *range* of X. That is,  $\mathscr{A} = \{X(\omega), \omega \in \Omega\}$ .

Definition: The expected value

**Definition:** The **expected value** (or mean, or expectation)

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable  $\boldsymbol{X}$  is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

#### Theorem:

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

$$E[X] = \sum_{a} a \times Pr[X = a]$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

$$E[X] = \sum_{a} a \times Pr[X = a]$$
$$= \sum_{a} a \times \sum_{\omega: X(\omega) = a} Pr[\omega]$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

$$E[X] = \sum_{a} a \times Pr[X = a]$$

$$= \sum_{a} a \times \sum_{\omega: X(\omega) = a} Pr[\omega]$$

$$= \sum_{a} \sum_{\omega: X(\omega) = a} X(\omega) Pr[\omega]$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

$$E[X] = \sum_{a} a \times Pr[X = a]$$

$$= \sum_{a} a \times \sum_{\omega: X(\omega) = a} Pr[\omega]$$

$$= \sum_{a} \sum_{\omega: X(\omega) = a} X(\omega) Pr[\omega]$$

$$= \sum_{\omega} X(\omega) Pr[\omega]$$

**Definition:** The **expected value** (or mean, or expectation) of a random variable X is

$$E[X] = \sum_{a} a \times Pr[X = a].$$

Theorem:

$$E[X] = \sum_{\omega} X(\omega) \times Pr[\omega].$$

$$E[X] = \sum_{a} a \times Pr[X = a]$$

$$= \sum_{a} a \times \sum_{\omega: X(\omega) = a} Pr[\omega]$$

$$= \sum_{a} \sum_{\omega: X(\omega) = a} X(\omega) Pr[\omega]$$

$$= \sum_{\alpha} X(\omega) Pr[\omega]$$

Flip a fair coin three times.

Flip a fair coin three times.

```
\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}.
```

Flip a fair coin three times.

 $\Omega = \{\textit{HHH}, \textit{HHT}, \textit{HTH}, \textit{THH}, \textit{HTT}, \textit{THT}, \textit{TTH}, \textit{TTT}\}.$ 

X = number of H's:  $\{3,2,2,2,1,1,1,0\}$ .

Flip a fair coin three times.

$$\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}.$$

$$X = \text{number of } H$$
's:  $\{3,2,2,2,1,1,1,0\}$ .

Thus,

$$\sum_{\omega} X(\omega) Pr[\omega] = \{3 + 2 + 2 + 2 + 1 + 1 + 1 + 0\} \times \frac{1}{8}.$$

Flip a fair coin three times.

$$\Omega = \{HHH, HHT, HTH, THH, HTT, THT, TTH, TTT\}.$$

$$X = \text{number of } H$$
's:  $\{3,2,2,2,1,1,1,0\}$ .

Thus,

$$\sum_{\omega} X(\omega) Pr[\omega] = \{3 + 2 + 2 + 2 + 1 + 1 + 1 + 0\} \times \frac{1}{8}.$$

Also,

$$\sum_{a} a \times Pr[X = a] = 3 \times \frac{1}{8} + 2 \times \frac{3}{8} + 1 \times \frac{3}{8} + 0 \times \frac{1}{8}.$$

Expected winnings for heads/tails games, with 3 flips?

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3, 1, 1, -1, 1, -1, -1, -3\}.$ 

$$E[X] = 3 \times \frac{1}{8}$$

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8}$$

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8}$$

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X: {HHH, HHT, HTH, HTH, THH, THT, TTH, TTT}  $\rightarrow$  {3,1,1,-1,1,-1,-1,-3}.

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

The expected value of *X* is not the value that you expect!

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable X:  $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3,1,1,-1,1,-1,-1,-3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

The expected value of X is not the value that you expect! It is the average value per experiment, if you perform the experiment many times:

### Win or Lose.

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable *X*:

 $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3, 1, 1, -1, 1, -1, -1, -3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

The expected value of *X* is not the value that you expect! It is the average value per experiment, if you perform the experiment many times:

$$\frac{X_1+\cdots+X_n}{n}$$
, when  $n\gg 1$ .

### Win or Lose.

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable *X*:

 $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3, 1, 1, -1, 1, -1, -1, -3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

The expected value of X is not the value that you expect! It is the average value per experiment, if you perform the experiment many times:

$$\frac{X_1+\cdots+X_n}{n}$$
, when  $n\gg 1$ .

The fact that this average converges to E[X] is a theorem:

### Win or Lose.

Expected winnings for heads/tails games, with 3 flips? Recall the definition of the random variable *X*:

 $\{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\} \rightarrow \{3, 1, 1, -1, 1, -1, -1, -3\}.$ 

$$E[X] = 3 \times \frac{1}{8} + 1 \times \frac{3}{8} - 1 \times \frac{3}{8} - 3 \times \frac{1}{8} = 0.$$

Can you ever win 0?

Apparently: expected value is not a common value, by any means.

The expected value of X is not the value that you expect! It is the average value per experiment, if you perform the experiment many times:

$$\frac{X_1+\cdots+X_n}{n}$$
, when  $n\gg 1$ .

The fact that this average converges to E[X] is a theorem: the Law of Large Numbers. (See later.)

# Multiple Random Variables.

Experiment: toss two coins.  $\Omega = \{HH, TH, HT, TT\}.$ 

# Multiple Random Variables.

Experiment: toss two coins.  $\Omega = \{HH, TH, HT, TT\}.$ 

$$X_1(\omega) = \left\{ egin{array}{ll} 1, & \mbox{if coin 1 is heads} \ 0, & \mbox{otherwise} \end{array} 
ight. \quad X_2(\omega) = \left\{ egin{array}{ll} 1, & \mbox{if coin 2 is heads} \ 0, & \mbox{otherwise} \end{array} 
ight.$$

**Joint Distribution:**  $\{(a, b, Pr[X = a, Y = b]) : a \in \mathcal{A}, b \in \mathcal{B}\}$ , where  $\mathcal{A}(\mathcal{B})$  is possible values of X(Y).

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathscr{A}, b \in \mathscr{B}\}$ , where  $\mathscr{A}(\mathscr{B})$  is possible values of X(Y).

$$\sum_{a \in \mathscr{A}, b \in \mathscr{B}} Pr[X = a, Y = b] =$$

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathscr{A}, b \in \mathscr{B}\}$ , where  $\mathscr{A}(\mathscr{B})$  is possible values of X(Y).

$$\sum_{a \in \mathscr{A}, b \in \mathscr{B}} Pr[X = a, Y = b] = 1$$

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathscr{A}, b \in \mathscr{B}\}$ , where  $\mathscr{A}(\mathscr{B})$  is possible values of X(Y).

$$\sum_{a \in \mathscr{A}, b \in \mathscr{B}} Pr[X = a, Y = b] = 1$$

Marginal for X:  $Pr[X = a] = \sum_{b \in \mathscr{B}} Pr[X = a, Y = b]$ . Marginal for Y:  $Pr[Y = b] = \sum_{a \in \mathscr{A}} Pr[X = a, Y = b]$ .

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathcal{A}, b \in \mathcal{B}\}$ , where  $\mathcal{A}(\mathcal{B})$  is possible values of X(Y).

$$\sum_{a \in \mathscr{A}, b \in \mathscr{B}} Pr[X = a, Y = b] = 1$$

Marginal for X:  $Pr[X = a] = \sum_{b \in \mathscr{B}} Pr[X = a, Y = b]$ . Marginal for Y:  $Pr[Y = b] = \sum_{a \in \mathscr{A}} Pr[X = a, Y = b]$ .

X/Y	1	2	3	Х
1	.2	.1	.1	.4
2	0	0	.3	.3
3	.1	0	.2	.3
Υ	.3	.1	.2	

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathcal{A}, b \in \mathcal{B}\}$ , where  $\mathcal{A}(\mathcal{B})$  is possible values of X(Y).

$$\sum_{a \in \mathscr{A}, b \in \mathscr{B}} Pr[X = a, Y = b] = 1$$

Marginal for X:  $Pr[X = a] = \sum_{b \in \mathscr{B}} Pr[X = a, Y = b]$ . Marginal for Y:  $Pr[Y = b] = \sum_{a \in \mathscr{A}} Pr[X = a, Y = b]$ .

X/Y	1	2	3	Χ
1	.2	.1	.1	.4
2	0	0	.3	.3
3	.1	0	.2	.3
Υ	.3	.1	.6	

**Joint Distribution:**  $\{(a,b,Pr[X=a,Y=b]): a \in \mathcal{A}, b \in \mathcal{B}\}$ , where  $\mathcal{A}(\mathcal{B})$  is possible values of X(Y).

$$\sum_{a\in\mathscr{A},b\in\mathscr{B}} Pr[X=a,Y=b]=1$$

Marginal for X:  $Pr[X = a] = \sum_{b \in \mathscr{B}} Pr[X = a, Y = b]$ . Marginal for Y:  $Pr[Y = b] = \sum_{a \in \mathscr{A}} Pr[X = a, Y = b]$ .

X/Y	1	2	3	Х
1	.2	.1	.1	.4
2	0	0	.3	.3
3	.1	0	.2	.3
Υ	.3	.1	.6	

Conditional Probability:  $Pr[X = a | Y = b] = \frac{Pr[X = a, Y = b]}{Pr[Y = b]}$ .

► Events A, B are independent if

▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- Events A, B, C are mutually independent if

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- Events A, B, C are mutually independent if A, B are independent, A, C are independent, B, C are independent

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- ► Events *A*, *B*, *C* are mutually independent if

A,B are independent, A,C are independent, B,C are independent

and  $Pr[A \cap B \cap C] = Pr[A]Pr[B]Pr[C]$ .

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- Events A, B, C are mutually independent if A, B are independent, A, C are independent, B, C are independent
  and Br[A ∩ B ∩ C] — Br[A] Br[B] Br[C]
  - and  $Pr[A \cap B \cap C] = Pr[A]Pr[B]Pr[C]$ .
- ▶ Events  $\{A_n, n \ge 0\}$  are mutually independent if ....

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- Events A, B, C are mutually independent if A, B are independent, A, C are independent, B, C are independent
  and Br[A ∩ B ∩ C] — Br[A] Br[B] Br[C]
  - and  $Pr[A \cap B \cap C] = Pr[A]Pr[B]Pr[C]$ .
- ▶ Events  $\{A_n, n \ge 0\}$  are mutually independent if ....
- ► Example:  $X, Y \in \{0,1\}$  two fair coin flips  $\Rightarrow X, Y, X \oplus Y$  are pairwise independent but not mutually independent.

- ▶ Events A, B are independent if  $Pr[A \cap B] = Pr[A]Pr[B]$ .
- ► Events A, B, C are mutually independent if
  A, B are independent, A, C are independent, B, C are independent
  - and  $Pr[A \cap B \cap C] = Pr[A]Pr[B]Pr[C]$ .
- ▶ Events  $\{A_n, n \ge 0\}$  are mutually independent if ....
- ► Example:  $X, Y \in \{0,1\}$  two fair coin flips  $\Rightarrow X, Y, X \oplus Y$  are pairwise independent but not mutually independent.
- Example:  $X, Y, Z \in \{0,1\}$  three fair coin flips are mutually independent.

**Definition:** Independence

**Definition:** Independence

The random variables *X* and *Y* are **independent** if and only if

$$Pr[Y = b|X = a] = Pr[Y = b]$$
, for all  $a$  and  $b$ .

**Definition:** Independence

The random variables *X* and *Y* are **independent** if and only if

$$Pr[Y = b|X = a] = Pr[Y = b]$$
, for all a and b.

Fact:

#### **Definition:** Independence

The random variables *X* and *Y* are **independent** if and only if

$$Pr[Y = b|X = a] = Pr[Y = b]$$
, for all a and b.

#### Fact:

X, Y are independent if and only if

$$Pr[X = a, Y = b] = Pr[X = a]Pr[Y = b]$$
, for all  $a$  and  $b$ .

#### **Definition:** Independence

The random variables *X* and *Y* are **independent** if and only if

$$Pr[Y = b|X = a] = Pr[Y = b]$$
, for all a and b.

#### Fact:

X, Y are independent if and only if

$$Pr[X = a, Y = b] = Pr[X = a]Pr[Y = b]$$
, for all  $a$  and  $b$ .

Follows from  $Pr[A \cap B] = Pr[A|B]Pr[B]$  (Product rule.)

### **Example 1**

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed:  $Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$ 

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

Indeed: 
$$Pr[X = 12, Y = 1] = 0 \neq Pr[X = 12]Pr[Y = 1] > 0$$
.

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

Indeed: 
$$Pr[X = 12, Y = 1] = 0 \neq Pr[X = 12]Pr[Y = 1] > 0$$
.

### Example 3

Flip a fair coin five times, X = number of Hs in first three flips, Y = number of Hs in last two flips. X and Y are independent.

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

Indeed: 
$$Pr[X = 12, Y = 1] = 0 \neq Pr[X = 12]Pr[Y = 1] > 0$$
.

#### Example 3

Flip a fair coin five times, X = number of Hs in first three flips, Y = number of Hs in last two flips. X and Y are independent.

Indeed:

$$Pr[X = a, Y = b] = {3 \choose a} {2 \choose b} 2^{-5}$$

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

Indeed: 
$$Pr[X = 12, Y = 1] = 0 \neq Pr[X = 12]Pr[Y = 1] > 0$$
.

#### Example 3

Flip a fair coin five times, X = number of Hs in first three flips, Y = number of Hs in last two flips. X and Y are independent.

Indeed:

$$Pr[X = a, Y = b] = {3 \choose a} {2 \choose b} 2^{-5} = {3 \choose a} 2^{-3} \times {2 \choose b} 2^{-2}$$

#### Example 1

Roll two die. X, Y = number of pips on the two dice. X, Y are independent.

Indeed: 
$$Pr[X = a, Y = b] = \frac{1}{36}, Pr[X = a] = Pr[Y = b] = \frac{1}{6}.$$

### Example 2

Roll two die. X = total number of pips, Y = number of pips on die 1 minus number on die 2. X and Y are not independent.

Indeed: 
$$Pr[X = 12, Y = 1] = 0 \neq Pr[X = 12]Pr[Y = 1] > 0$$
.

### Example 3

Flip a fair coin five times, X = number of Hs in first three flips, Y = number of Hs in last two flips. X and Y are independent.

Indeed:

$$Pr[X = a, Y = b] = {3 \choose a} {2 \choose b} 2^{-5} = {3 \choose a} 2^{-3} \times {2 \choose b} 2^{-2} = Pr[X = a] Pr[Y = b].$$

## Linearity of Expectation

Theorem:

$$E[X+Y] = E[X] + E[Y]$$

```
Theorem:

E[X + Y] = E[X] + E[Y]

E[cX] = cE[X]
```

#### Theorem:

$$E[X + Y] = E[X] + E[Y]$$
$$E[cX] = cE[X]$$

Proof:  $E[X] = \sum_{\omega \in \Omega} X(\omega) \times P[\omega]$ .

#### Theorem:

$$E[X + Y] = E[X] + E[Y]$$
$$E[cX] = cE[X]$$

Proof:  $E[X] = \sum_{\omega \in \Omega} X(\omega) \times P[\omega]$ .

$$E[X + Y] = \sum_{\omega \in \Omega} (X(\omega) + Y(\omega))Pr[\omega]$$

$$= \sum_{\omega \in \Omega} X(\omega)Pr[\omega] + Y(\omega)Pr[\omega]$$

$$= \sum_{\omega \in \Omega} X(\omega)Pr[\omega] + \sum_{\omega \in \Omega} Y(\omega)Pr[\omega]$$

$$= E[X] + E[Y]$$

**Definition** 

#### **Definition**

Let A be an event. The random variable X defined by

#### **Definition**

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

#### **Definition**

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

Note that 
$$Pr[X = 1] =$$

#### **Definition**

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

Note that 
$$Pr[X = 1] = Pr[A]$$
 and  $Pr[X = 0] =$ 

#### **Definition**

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

Note that 
$$Pr[X = 1] = Pr[A]$$
 and  $Pr[X = 0] = 1 - Pr[A]$ .

#### Definition

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

is called the indicator of the event A.

Note that Pr[X = 1] = Pr[A] and Pr[X = 0] = 1 - Pr[A].

Hence,

$$E[X] = 1 \times Pr[X = 1] + 0 \times Pr[X = 0] = Pr[A].$$

#### Definition

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

is called the indicator of the event A.

Note that Pr[X = 1] = Pr[A] and Pr[X = 0] = 1 - Pr[A].

Hence,

$$E[X] = 1 \times Pr[X = 1] + 0 \times Pr[X = 0] = Pr[A].$$

This random variable  $X(\omega)$  is sometimes written as

$$1\{\omega \in A\}$$
 or  $1_A(\omega)$ .

#### Definition

Let A be an event. The random variable X defined by

$$X(\omega) = \begin{cases} 1, & \text{if } \omega \in A \\ 0, & \text{if } \omega \notin A \end{cases}$$

is called the indicator of the event A.

Note that Pr[X = 1] = Pr[A] and Pr[X = 0] = 1 - Pr[A].

Hence,

$$E[X] = 1 \times Pr[X = 1] + 0 \times Pr[X = 0] = Pr[A].$$

This random variable  $X(\omega)$  is sometimes written as

$$1\{\omega \in A\}$$
 or  $1_A(\omega)$ .

Thus, we will write  $X = 1_A$ .

Theorem:

Theorem: Expectation is linear

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

$$E[a_1X_1+\cdots+a_nX_n]$$

**Theorem:** Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

$$E[a_1X_1 + \cdots + a_nX_n]$$
  
=  $\sum_{\omega} (a_1X_1 + \cdots + a_nX_n)(\omega)Pr[\omega]$ 

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

$$E[a_1X_1 + \dots + a_nX_n]$$

$$= \sum_{\omega} (a_1X_1 + \dots + a_nX_n)(\omega)Pr[\omega]$$

$$= \sum_{\omega} (a_1X_1(\omega) + \dots + a_nX_n(\omega))Pr[\omega]$$

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

$$E[a_1X_1 + \dots + a_nX_n]$$

$$= \sum_{\omega} (a_1X_1 + \dots + a_nX_n)(\omega)Pr[\omega]$$

$$= \sum_{\omega} (a_1X_1(\omega) + \dots + a_nX_n(\omega))Pr[\omega]$$

$$= a_1\sum_{\omega} X_1(\omega)Pr[\omega] + \dots + a_n\sum_{\omega} X_n(\omega)Pr[\omega]$$

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

$$E[a_1X_1 + \dots + a_nX_n]$$

$$= \sum_{\omega} (a_1X_1 + \dots + a_nX_n)(\omega)Pr[\omega]$$

$$= \sum_{\omega} (a_1X_1(\omega) + \dots + a_nX_n(\omega))Pr[\omega]$$

$$= a_1\sum_{\omega} X_1(\omega)Pr[\omega] + \dots + a_n\sum_{\omega} X_n(\omega)Pr[\omega]$$

$$= a_1E[X_1] + \dots + a_nE[X_n].$$

Theorem: Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

Proof:

$$E[a_1X_1 + \dots + a_nX_n]$$

$$= \sum_{\omega} (a_1X_1 + \dots + a_nX_n)(\omega)Pr[\omega]$$

$$= \sum_{\omega} (a_1X_1(\omega) + \dots + a_nX_n(\omega))Pr[\omega]$$

$$= a_1\sum_{\omega} X_1(\omega)Pr[\omega] + \dots + a_n\sum_{\omega} X_n(\omega)Pr[\omega]$$

$$= a_1E[X_1] + \dots + a_nE[X_n].$$

Note:

**Theorem:** Expectation is linear

$$E[a_1X_1 + \cdots + a_nX_n] = a_1E[X_1] + \cdots + a_nE[X_n].$$

Proof:

$$E[a_1X_1 + \dots + a_nX_n]$$

$$= \sum_{\omega} (a_1X_1 + \dots + a_nX_n)(\omega)Pr[\omega]$$

$$= \sum_{\omega} (a_1X_1(\omega) + \dots + a_nX_n(\omega))Pr[\omega]$$

$$= a_1\sum_{\omega} X_1(\omega)Pr[\omega] + \dots + a_n\sum_{\omega} X_n(\omega)Pr[\omega]$$

$$= a_1E[X_1] + \dots + a_nE[X_n].$$

Note: If we had defined  $Y = a_1 X_1 + \cdots + a_n X_n$  has had tried to compute  $E[Y] = \sum_y y Pr[Y = y]$ , we would have been in trouble!

Roll a die n times.

Roll a die n times.

Roll a die n times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

Roll a die *n* times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \cdots + X_n]$$

Roll a die n times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n],$ 

Roll a die n times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity

Roll a die n times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ ,

Roll a die n times.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Roll a die n times.

 $X_m$  = number of pips on roll m.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} =$$

Roll a die n times.

 $X_m$  = number of pips on roll m.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} = \frac{6 \times 7}{2} \times \frac{1}{6} =$$

Roll a die n times.

 $X_m$  = number of pips on roll m.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} = \frac{6 \times 7}{2} \times \frac{1}{6} = \frac{7}{2}.$$

Roll a die n times.

 $X_m$  = number of pips on roll m.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} = \frac{6 \times 7}{2} \times \frac{1}{6} = \frac{7}{2}.$$

Hence,

$$E[X]=\frac{7n}{2}$$
.

## Using Linearity - 1: Pips (dots) on dice

Roll a die n times.

 $X_m$  = number of pips on roll m.

$$X = X_1 + \cdots + X_n$$
 = total number of pips in  $n$  rolls.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because the  $X_m$  have the same distribution

Now,

$$E[X_1] = 1 \times \frac{1}{6} + \dots + 6 \times \frac{1}{6} = \frac{6 \times 7}{2} \times \frac{1}{6} = \frac{7}{2}.$$

Hence,

$$E[X]=\frac{7n}{2}.$$

Note: Computing  $\sum_{x} xPr[X = x]$  directly is not easy!

Hand out assignments at random to *n* students.

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

 $X = X_1 + \cdots + X_n$  where

 $X_m = 1$  {student m gets his/her own assignment back}.

Hand out assignments at random to n students.

X = number of students that get their own assignment back.

 $X = X_1 + \cdots + X_n$  where

 $X_m = 1$ {student m gets his/her own assignment back}.

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \cdots + X_n]$$

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n],$ 

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$ {student  $m$  gets his/her own assignment back}.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ ,

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because all the  $X_m$  have the same distribution

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \dots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \dots + X_n]$$
  
=  $E[X_1] + \dots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because all the  $X_m$  have the same distribution  
=  $nPr[X_1 = 1]$ ,

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because all the  $X_m$  have the same distribution  
=  $nPr[X_1 = 1]$ , because  $X_1$  is an indicator

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$ {student  $m$  gets his/her own assignment back}.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because all the  $X_m$  have the same distribution  
=  $nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
=  $n(1/n)$ ,

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where

 $X_m = 1$ {student m gets his/her own assignment back}.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
=  $E[X_1] + \cdots + E[X_n]$ , by linearity  
=  $nE[X_1]$ , because all the  $X_m$  have the same distribution  
=  $nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
=  $n(1/n)$ , because student 1 is equally likely  
to get any one of the  $n$  assignments

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$ {student  $m$  gets his/her own assignment back}.

$$E[X] = E[X_1 + \cdots + X_n]$$
  
 $= E[X_1] + \cdots + E[X_n]$ , by linearity  
 $= nE[X_1]$ , because all the  $X_m$  have the same distribution  
 $= nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
 $= n(1/n)$ , because student 1 is equally likely  
to get any one of the  $n$  assignments  
 $= 1$ .

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$  {student  $m$  gets his/her own assignment back}.

One has

$$E[X] = E[X_1 + \cdots + X_n]$$
  
 $= E[X_1] + \cdots + E[X_n]$ , by linearity  
 $= nE[X_1]$ , because all the  $X_m$  have the same distribution  
 $= nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
 $= n(1/n)$ , because student 1 is equally likely  
to get any one of the  $n$  assignments  
 $= 1$ .

Note that linearity holds even though the  $X_m$  are not independent (whatever that means).

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$  {student  $m$  gets his/her own assignment back}.

One has

$$E[X] = E[X_1 + \cdots + X_n]$$
  
 $= E[X_1] + \cdots + E[X_n]$ , by linearity  
 $= nE[X_1]$ , because all the  $X_m$  have the same distribution  
 $= nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
 $= n(1/n)$ , because student 1 is equally likely  
to get any one of the  $n$  assignments  
 $= 1$ .

Note that linearity holds even though the  $X_m$  are not independent (whatever that means).

Note: What is Pr[X = m]?

Hand out assignments at random to *n* students.

X = number of students that get their own assignment back.

$$X = X_1 + \cdots + X_n$$
 where  $X_m = 1$  {student  $m$  gets his/her own assignment back}.

One has

$$E[X] = E[X_1 + \cdots + X_n]$$
  
 $= E[X_1] + \cdots + E[X_n]$ , by linearity  
 $= nE[X_1]$ , because all the  $X_m$  have the same distribution  
 $= nPr[X_1 = 1]$ , because  $X_1$  is an indicator  
 $= n(1/n)$ , because student 1 is equally likely  
to get any one of the  $n$  assignments  
 $= 1$ .

Note that linearity holds even though the  $X_m$  are not independent (whatever that means).

Note: What is Pr[X = m]? Tricky ....

Flip n coins with heads probability p.

Flip n coins with heads probability p. X - number of heads

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

E[X]

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i]$$

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ...

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ... Or...

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

$$X_i = \begin{cases} 1 & \text{if } i \text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

Flip *n* coins with heads probability *p*. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

$$X_i = \begin{cases} 1 & \text{if } i \text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"]$$

Flip *n* coins with heads probability *p*. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

$$X_i = \begin{cases} 1 & \text{if } i \text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"] = p.$$

Flip *n* coins with heads probability *p*. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ... Or... a better approach: Let

$$X_i = \begin{cases} 1 & \text{if } i \text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"] = p.$$

Flip *n* coins with heads probability *p*. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ... Or... a better approach: Let

$$X_i = \begin{cases} 1 & \text{if } i \text{th flip is heads} \\ 0 & \text{otherwise} \end{cases}$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"] = p.$$

$$E[X] = E[X_1] + E[X_2] + \cdots + E[X_n]$$

Flip *n* coins with heads probability *p*. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ... Or... a better approach: Let

$$X_i = \left\{ egin{array}{ll} 1 & ext{ if } i ext{th flip is heads} \\ 0 & ext{ otherwise} \end{array} 
ight.$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"] = p.$$

$$E[X] = E[X_1] + E[X_2] + \cdots + E[X_n] = n \times E[X_i]$$

Flip n coins with heads probability p. X - number of heads Binomial Distibution: Pr[X = i], for each i.

$$Pr[X=i] = \binom{n}{i} p^{i} (1-p)^{n-i}.$$

$$E[X] = \sum_{i} i \times Pr[X = i] = \sum_{i} i \times \binom{n}{i} p^{i} (1 - p)^{n - i}.$$

Uh oh. ... Or... a better approach: Let

$$X_i = \left\{ egin{array}{ll} 1 & ext{ if } i ext{th flip is heads} \\ 0 & ext{ otherwise} \end{array} 
ight.$$

$$E[X_i] = 1 \times Pr["heads"] + 0 \times Pr["tails"] = p.$$

$$E[X] = E[X_1] + E[X_2] + \cdots + E[X_n] = n \times E[X_i] = np.$$

# Using Linearity - 4

## Using Linearity - 4

Assume  $\boldsymbol{A}$  and  $\boldsymbol{B}$  are disjoint events.

Assume *A* and *B* are disjoint events. Then  $1_{A \cup B}(\omega) = 1_A(\omega) + 1_B(\omega)$ .

$$\textit{Pr}[A \cup B] = E[1_{A \cup B}]$$

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] =$$

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] =$$

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] = Pr[A] + Pr[B].$$

Assume A and B are disjoint events. Then  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)$ . Taking expectation, we get

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] = Pr[A] + Pr[B].$$

In general,  $1_{A \cup B}(\omega) = 1_A(\omega) + 1_B(\omega) - 1_{A \cap B}(\omega)$ .

Assume A and B are disjoint events. Then  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)$ . Taking expectation, we get

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] = Pr[A] + Pr[B].$$

In general,  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)-1_{A\cap B}(\omega)$ . Taking expectation, we get  $Pr[A\cup B]=Pr[A]+Pr[B]-Pr[A\cap B]$ .

Assume A and B are disjoint events. Then  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)$ . Taking expectation, we get

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] = Pr[A] + Pr[B].$$

In general,  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)-1_{A\cap B}(\omega)$ . Taking expectation, we get  $Pr[A\cup B]=Pr[A]+Pr[B]-Pr[A\cap B]$ .

Observe that if  $Y(\omega) = b$  for all  $\omega$ , then E[Y] = b.

Assume A and B are disjoint events. Then  $1_{A\cup B}(\omega)=1_A(\omega)+1_B(\omega)$ . Taking expectation, we get

$$Pr[A \cup B] = E[1_{A \cup B}] = E[1_A + 1_B] = E[1_A] + E[1_B] = Pr[A] + Pr[B].$$

In general,  $1_{A \cup B}(\omega) = 1_A(\omega) + 1_B(\omega) - 1_{A \cap B}(\omega)$ . Taking expectation, we get  $Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$ .

Observe that if  $Y(\omega) = b$  for all  $\omega$ , then E[Y] = b. Thus, E[X + b] = E[X] + b.

Experiment: Throw m balls into n bins.

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Experiment: Throw *m* balls into *n* bins.

*Y* - number of empty bins.

Distribution is horrible.

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

Expectation?

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

$$Y = X_1 + \cdots X_n$$
.

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

$$Y = X_1 + \cdots X_n$$
.

$$Pr[X_1 = 1] = (1 - \frac{1}{n})^m$$
.

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

$$Y = X_1 + \cdots X_n$$
.

$$Pr[X_1 = 1] = (1 - \frac{1}{n})^m$$
.  $\to E[Y] = n(1 - \frac{1}{n})^m$ .

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

Expectation?  $X_i$  - indicator for bin i being empty.

$$Y = X_1 + \cdots X_n$$
.

$$Pr[X_1 = 1] = (1 - \frac{1}{n})^m$$
.  $\rightarrow E[Y] = n(1 - \frac{1}{n})^m$ .

For n = m and large n,  $(1 - 1/n)^n \approx \frac{1}{e}$ .

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

Expectation?  $X_i$  - indicator for bin i being empty.

$$Y = X_1 + \cdots X_n$$
.

$$Pr[X_1 = 1] = (1 - \frac{1}{n})^m$$
.  $\rightarrow E[Y] = n(1 - \frac{1}{n})^m$ .

For n = m and large n,  $(1 - 1/n)^n \approx \frac{1}{e}$ .

Experiment: Throw *m* balls into *n* bins.

Y - number of empty bins.

Distribution is horrible.

$$Y = X_1 + \cdots X_n$$
.

$$Pr[X_1 = 1] = (1 - \frac{1}{n})^m$$
.  $\to E[Y] = n(1 - \frac{1}{n})^m$ .

For 
$$n = m$$
 and large  $n$ ,  $(1 - 1/n)^n \approx \frac{1}{e}$ .

 $<sup>\</sup>frac{n}{e} \approx 0.368n$  empty bins on average.

**Experiment:** Get random coupon from n until get all n coupons.

**Experiment:** Get random coupon from n until get all n coupons.

**Outcomes:** {123145...,56765...}

**Experiment:** Get random coupon from *n* until get all *n* coupons.

**Outcomes:** {123145...,56765...}

**Random Variable:** *X* - length of outcome.

**Experiment:** Get random coupon from *n* until get all *n* coupons.

**Outcomes:** {123145...,56765...}

**Random Variable:** *X* - length of outcome.

**Experiment:** Get random coupon from n until get all n coupons.

**Outcomes:** {123145...,56765...}

**Random Variable:** *X* - length of outcome.

Today: E[X]?

**Experiment:** Get random coupon from n until get all n coupons.

**Outcomes:** {123145...,56765...}

**Random Variable:** *X* - length of outcome.

Today: E[X]?

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1}p, n \ge 1$ .

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1}p$ ,  $n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1-p)^{n-1}p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p+2(1-p)p+3(1-p)^2p+4(1-p)^3p+\cdots$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1} p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p+2(1-p)p+3(1-p)^2p+4(1-p)^3p+\cdots$$

$$(1-p)E[X] = (1-p)p+2(1-p)^2p+3(1-p)^3p+\cdots$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1} p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p+2(1-p)p+3(1-p)^{2}p+4(1-p)^{3}p+\cdots$$

$$(1-p)E[X] = (1-p)p+2(1-p)^{2}p+3(1-p)^{3}p+\cdots$$

$$pE[X] = p+(1-p)p+(1-p)^{2}p+(1-p)^{3}p+\cdots$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1} p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p + 2(1-p)p + 3(1-p)^2p + 4(1-p)^3p + \cdots$$

$$(1-p)E[X] = (1-p)p + 2(1-p)^2p + 3(1-p)^3p + \cdots$$

$$pE[X] = p + (1-p)p + (1-p)^2p + (1-p)^3p + \cdots$$
by subtracting the previous two identities

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1}p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p+2(1-p)p+3(1-p)^{2}p+4(1-p)^{3}p+\cdots$$

$$(1-p)E[X] = (1-p)p+2(1-p)^{2}p+3(1-p)^{3}p+\cdots$$

$$pE[X] = p+(1-p)p+(1-p)^{2}p+(1-p)^{3}p+\cdots$$
by subtracting the previous two identities
$$= \sum_{n=0}^{\infty} Pr[X=n] =$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1} p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

$$E[X] = p + 2(1-p)p + 3(1-p)^2p + 4(1-p)^3p + \cdots$$

$$(1-p)E[X] = (1-p)p + 2(1-p)^2p + 3(1-p)^3p + \cdots$$

$$pE[X] = p + (1-p)p + (1-p)^2p + (1-p)^3p + \cdots$$
by subtracting the previous two identities
$$= \sum_{n=1}^{\infty} Pr[X=n] = 1.$$

$$X =_D G(p)$$
, i.e.,  $Pr[X = n] = (1 - p)^{n-1}p, n \ge 1$ .

One has

$$E[X] = \sum_{n=1}^{\infty} nPr[X = n] = \sum_{n=1}^{\infty} n(1-p)^{n-1}p.$$

Thus,

$$E[X] = p+2(1-p)p+3(1-p)^{2}p+4(1-p)^{3}p+\cdots$$

$$(1-p)E[X] = (1-p)p+2(1-p)^{2}p+3(1-p)^{3}p+\cdots$$

$$pE[X] = p+(1-p)p+(1-p)^{2}p+(1-p)^{3}p+\cdots$$
by subtracting the previous two identities

$$= \sum_{n=1}^{\infty} Pr[X=n] = 1.$$

Hence,

$$E[X]=\frac{1}{p}$$
.

*X*-time to get *n* coupons.

*X*-time to get *n* coupons.

 $X_1$  - time to get first coupon.

*X*-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .

*X*-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

Pr["get second coupon"|"got milk "

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"|"got milk first coupon"] = \frac{n-1}{n}$ 

 $E[X_2]$ ?

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

 $E[X_2]$ ? Geometric

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

 $E[X_2]$ ? Geometric!

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"|"got milk first coupon"] = \frac{n-1}{n}$ 

 $E[X_2]$ ? Geometric!!

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

 $E[X_2]$ ? Geometric!!!

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr[\text{"get second coupon"}|\text{"got milk first coupon"}] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\Longrightarrow E[X_2] = \frac{1}{\rho} =$ 

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\Longrightarrow E[X_2] = \frac{1}{\rho} = \frac{1}{\frac{n-1}{n}}$ 

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{\rho} = \frac{1}{\frac{n-1}{n}} = \frac{n}{n-1}$ .

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{p}} = \frac{n}{n-1}$ .

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$Pr["getting ith coupon|"got i - 1rst coupons"] = \frac{n - (i - 1)}{n} = \frac{n - i + 1}{n}$$

 $E[X_i]$ 

X-time to get n coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$Pr["getting ith coupon|"got i - 1rst coupons"] = \frac{n - (i - 1)}{n} = \frac{n - i + 1}{n}$$

$$E[X_i] = \frac{1}{p}$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$Pr["getting ith coupon|"got i-1 rst coupons"] = \frac{n-(i-1)}{n} = \frac{n-i+1}{n}$$

$$E[X_i] = \frac{1}{\rho} = \frac{n}{n-i+1},$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$E[X_i] = \frac{1}{\rho} = \frac{n}{n-i+1}, i = 1, 2, ..., n.$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, ..., n.$$

$$E[X] = E[X_1] + \cdots + E[X_n] =$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{2}} = \frac{n}{n-1}$ .

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, ..., n.$$

$$E[X] = E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1}$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{\rho} = \frac{1}{\frac{n-1}{\rho}} = \frac{n}{n-1}$ .

$$E[X_i] = \frac{1}{p} = \frac{n}{n-i+1}, i = 1, 2, ..., n.$$

$$E[X] = E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1}$$
$$= n(1 + \frac{1}{2} + \dots + \frac{1}{n}) =: nH(n)$$

X-time to get *n* coupons.

 $X_1$  - time to get first coupon. Note:  $X_1 = 1$ .  $E(X_1) = 1$ .

 $X_2$  - time to get second coupon after getting first.

 $Pr["get second coupon"]"got milk first coupon"] = \frac{n-1}{n}$ 

$$E[X_2]$$
? Geometric!!!  $\implies E[X_2] = \frac{1}{p} = \frac{1}{\frac{n-1}{p}} = \frac{n}{n-1}$ .

$$E[X_i] = \frac{1}{\rho} = \frac{n}{n-i+1}, i = 1, 2, ..., n.$$

$$E[X] = E[X_1] + \dots + E[X_n] = \frac{n}{n} + \frac{n}{n-1} + \frac{n}{n-2} + \dots + \frac{n}{1}$$
$$= n(1 + \frac{1}{2} + \dots + \frac{1}{n}) =: nH(n) \approx n(\ln n + \gamma)$$

#### Review: Harmonic sum

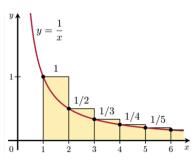
$$H(n) = 1 + \frac{1}{2} + \cdots + \frac{1}{n} \approx \int_{1}^{n} \frac{1}{x} dx = \ln(n).$$

.

#### Review: Harmonic sum

$$H(n) = 1 + \frac{1}{2} + \dots + \frac{1}{n} \approx \int_{1}^{n} \frac{1}{x} dx = \ln(n).$$

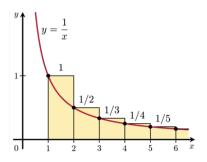
•



#### Review: Harmonic sum

$$H(n) = 1 + \frac{1}{2} + \cdots + \frac{1}{n} \approx \int_{1}^{n} \frac{1}{x} dx = \ln(n).$$

•

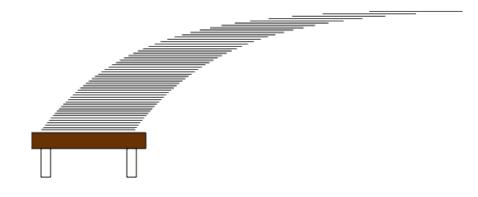


#### A good approximation is

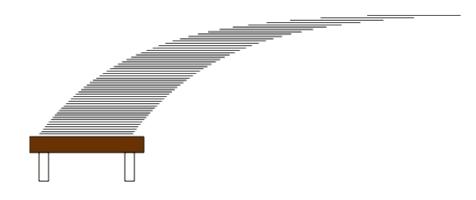
 $H(n) \approx \ln(n) + \gamma$  where  $\gamma \approx 0.58$  (Euler-Mascheroni constant).

Consider this stack of cards (no glue!):

Consider this stack of cards (no glue!):

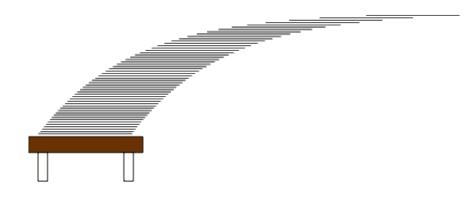


Consider this stack of cards (no glue!):



If each card has length 2, the stack can extend H(n) to the right of the table.

Consider this stack of cards (no glue!):



If each card has length 2, the stack can extend H(n) to the right of the table. As n increases, you can go as far as you want!

#### **Paradox**

# par·a·dox

/'perə däks/

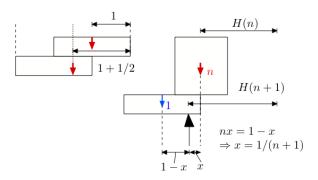
noun

a statement or proposition that, despite sound (or apparently sound) reasoning from acceptable premises, leads to a conclusion that seems senseless, logically unacceptable, or self-contradictory.

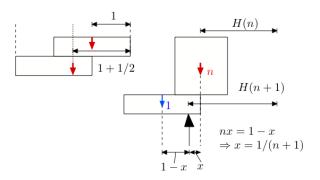
"a potentially serious conflict between quantum mechanics and the general theory of relativity known as the information paradox"

- a seemingly absurd or self-contradictory statement or proposition that when investigated or explained may prove to be well founded or true.
   "in a paradox, he has discovered that stepping back from his job has increased the rewards he gleans from it"
   synonyms: contradiction, contradiction in terms, self-contradiction, inconsistency, incongruity; More
- a situation, person, or thing that combines contradictory features or qualities.
   "the mingling of deciduous trees with elements of desert flora forms a fascinating ecological paradox"

## Stacking

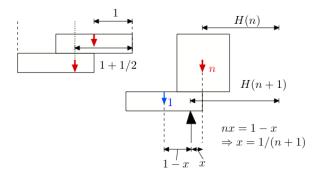


## Stacking



The cards have width 2.

## Stacking



The cards have width 2. Induction shows that the center of gravity after n cards is H(n) away from the right-most edge. Video.

# Calculating E[g(X)]

## Calculating E[g(X)]Let Y = g(X).

Let Y = g(X). Assume that we know the distribution of X.

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

Method 1:

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)]$$
 where  $g^{-1}(x) = \{x \in \Re : g(x) = y\}.$ 

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)]$$
 where  $g^{-1}(x) = \{x \in \Re : g(x) = y\}.$ 

This is typically rather tedious!

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

**Method 2:** We use the following result.

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

$$E[g(X)] = \sum_{\omega} g(X(\omega)) Pr[\omega]$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

$$E[g(X)] = \sum_{\omega} g(X(\omega))Pr[\omega] = \sum_{X} \sum_{\omega \in X^{-1}(X)} g(X(\omega))Pr[\omega]$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

$$E[g(X)] = \sum_{\omega} g(X(\omega)) Pr[\omega] = \sum_{x} \sum_{\omega \in X^{-1}(x)} g(X(\omega)) Pr[\omega]$$
$$= \sum_{x} \sum_{\omega \in X^{-1}(x)} g(x) Pr[\omega]$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)] \text{ where } g^{-1}(x) = \{x \in \Re : g(x) = y\}.$$

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

$$E[g(X)] = \sum_{\omega} g(X(\omega))Pr[\omega] = \sum_{x} \sum_{\omega \in X^{-1}(x)} g(X(\omega))Pr[\omega]$$
$$= \sum_{x} \sum_{\omega \in X^{-1}(x)} g(x)Pr[\omega] = \sum_{x} g(x) \sum_{\omega \in X^{-1}(x)} Pr[\omega]$$

Let Y = g(X). Assume that we know the distribution of X.

We want to calculate E[Y].

**Method 1:** We calculate the distribution of *Y*:

$$Pr[Y = y] = Pr[X \in g^{-1}(y)]$$
 where  $g^{-1}(x) = \{x \in \Re : g(x) = y\}.$ 

This is typically rather tedious!

Method 2: We use the following result.

Theorem:

$$E[g(X)] = \sum_{x} g(x) Pr[X = x].$$

$$E[g(X)] = \sum_{\omega} g(X(\omega))Pr[\omega] = \sum_{x} \sum_{\omega \in X^{-1}(x)} g(X(\omega))Pr[\omega]$$
$$= \sum_{x} \sum_{\omega \in X^{-1}(x)} g(x)Pr[\omega] = \sum_{x} g(x) \sum_{\omega \in X^{-1}(x)} Pr[\omega]$$
$$= \sum_{x} g(x)Pr[X = x].$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ .

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^2 \frac{1}{6}$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

Let *X* be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

Let *X* be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

$$Y = \begin{cases} 4, & \text{w.p. } \frac{2}{6} \end{cases}$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

$$Y = \begin{cases} 4, & \text{w.p. } \frac{2}{6} \\ 1, & \text{w.p. } \frac{2}{6} \end{cases}$$

Let *X* be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

$$Y = \begin{cases} 4, & \text{w.p. } \frac{2}{6} \\ 1, & \text{w.p. } \frac{2}{6} \\ 0, & \text{w.p. } \frac{1}{6} \end{cases}$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

$$Y = \left\{ \begin{array}{ll} 4, & \text{w.p. } \frac{2}{6} \\ 1, & \text{w.p. } \frac{2}{6} \\ 0, & \text{w.p. } \frac{1}{6} \\ 9, & \text{w.p. } \frac{1}{6}. \end{array} \right.$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

$$Y = \begin{cases} 4, & \text{w.p. } \frac{2}{6} \\ 1, & \text{w.p. } \frac{2}{6} \\ 0, & \text{w.p. } \frac{1}{6} \\ 9, & \text{w.p. } \frac{1}{6} \end{cases}$$

Let X be uniform in  $\{-2, -1, 0, 1, 2, 3\}$ .

Let also  $g(X) = X^2$ . Then (method 2)

$$E[g(X)] = \sum_{x=-2}^{3} x^{2} \frac{1}{6}$$

$$= \{4+1+0+1+4+9\} \frac{1}{6} = \frac{19}{6}.$$

Method 1 - We find the distribution of  $Y = X^2$ :

$$Y = \begin{cases} 4, & \text{w.p. } \frac{2}{6} \\ 1, & \text{w.p. } \frac{2}{6} \\ 0, & \text{w.p. } \frac{1}{6} \\ 9, & \text{w.p. } \frac{1}{6} \end{cases}$$

Thus,

$$E[Y] = 4\frac{2}{6} + 1\frac{2}{6} + 0\frac{1}{6} + 9\frac{1}{6} = \frac{19}{6}.$$

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Probability Space:  $\Omega,$   $Pr[\omega] \geq 0,$   $\sum_{\omega} Pr[\omega] = 1.$ 

Random Variable: Function on Sample Space.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a]$ 

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_{a} Pr[X = a] = 1$ .

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_{a} Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_{a} Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

$$\sum_{a,b} Pr[X=a,Y=b]$$

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

 $\sum_{a,b} Pr[X=a,Y=b]=1.$ 

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

 $\sum_{a,b} Pr[X=a,Y=b]=1.$ 

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

 $\sum_{a,b} Pr[X=a,Y=b]=1.$ 

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

 $\sum_{a,b} Pr[X=a,Y=b]=1.$ 

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Indicators: Empty bins, Fixed points.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

$$\sum_{a,b} Pr[X=a,Y=b] = 1.$$

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Indicators: Empty bins, Fixed points.

Time to Coupon: Sum times to "next" coupon.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

$$\sum_{a,b} Pr[X=a,Y=b] = 1.$$

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Indicators: Empty bins, Fixed points.

Time to Coupon: Sum times to "next" coupon.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

$$\sum_{a,b} Pr[X=a,Y=b]=1.$$

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Indicators: Empty bins, Fixed points.

Time to Coupon: Sum times to "next" coupon.

Y = f(X) is Random Variable.

Probability Space:  $\Omega$ ,  $Pr[\omega] \ge 0$ ,  $\sum_{\omega} Pr[\omega] = 1$ .

Random Variable: Function on Sample Space.

Distribution: Function  $Pr[X = a] \ge 0$ .  $\sum_a Pr[X = a] = 1$ .

Expectation:  $E[X] = \sum_{\omega} Pr[\omega] = \sum_{a} Pr[X = a]$ .

Many Random Variables: each one function on a sample space.

Joint Distributions: Function  $Pr[X = a, Y = b] \ge 0$ .

$$\sum_{a,b} Pr[X=a,Y=b]=1.$$

Linearity of Expectation: E[X + Y] = E[X] + E[Y].

Applications: compute expectations by decomposing.

Indicators: Empty bins, Fixed points.

Time to Coupon: Sum times to "next" coupon.

Y = f(X) is Random Variable.

Distribution of *Y* from distribution of *X*.